

Pacific Journal of Mathematics

**SAMPLE FUNCTION BEHAVIOR OF INCREASING
PROCESSES WITH STATIONARY, INDEPENDENT
INCREMENTS**

BERT E. FRISTEDT

SAMPLE FUNCTION BEHAVIOR OF INCREASING PROCESSES WITH STATIONARY, INDEPENDENT INCREMENTS

BERT E. FRISTEDT

In this paper we are concerned with the sample functions of increasing stochastic processes, X_ν , having stationary, independent increments; normalized so that X_ν has no deterministic linear component and $X_\nu(0) = 0$, (i.e., X_ν is a subordinator).

For h a fixed function, we are interested in the following two events:

$$\begin{aligned} &\{\omega: X_\nu(t, \omega) > h(t) \text{ infinitely often as } t \rightarrow 0\}, \\ &\{\omega: X_\nu(t, \omega) > h(t) \text{ infinitely often as } t \rightarrow \infty\}. \end{aligned}$$

In case X_ν is a stable process, Khinchin has given integral tests to apply to a wide class of h 's in order to decide whether one, the other, or both of these two events have probability zero or one. The purpose of this paper is to give similar results, without assuming X_ν to be stable.

We also prove (Theorem 3) a variation-type theorem concerning the sample functions. Theorem 4 is an L_1 convergence theorem for the distribution function as time goes to zero.

2. Notation. We let ν be a measure on $(0, \infty)$ with

$$\int_0^\infty y(1+y)^{-1} \nu(dy) < \infty.$$

Then, we let

$$\begin{aligned} g_\nu(u) &= \int_0^\infty (1 - e^{-uy}) \nu(dy); \\ \varphi_\nu(t, u) &= \exp(-tg_\nu(u)). \end{aligned}$$

We let X_ν be a function of two variables; the first variable being an element in $[0, \infty)$ and the second variable being an element in a probability space Ω with probability measure P . We take $X_\nu(0, \omega) = 0$ for all $\omega \in \Omega$, and we take $X_\nu(\cdot, \omega)$ to be an increasing right continuous function (The range of X_ν is taken to be a subset of $[0, \infty)$). We require that

$$\int_0^\infty e^{-ux} d_2 F_\nu(t, x) = \varphi_\nu(t, u)$$

where

$$F_\nu(t, x) = P\{\omega: X_\nu(t, \omega) < x\}.$$

It follows ([1] and [6], 417-424) that X_ν is an increasing stochastic process having stationary, independent increments. Moreover, if $Y(t, \omega) = X_\nu(t, \omega) - ct$ where $c > 0$, then Y will not be an increasing process (i.e. X_ν has no deterministic linear component.). It is true ([1] and [6], 417-424) that every increasing stochastic process having stationary, independent increments and having the property mentioned in the preceding sentence—whose value at 0 is 0 and which is right continuous—is an X_ν for an appropriate choice of ν . The measure ν is called the Levy measure corresponding to the process X_ν .

If we remove an appropriate set of measure zero from Ω , we can make some further interesting statements ([4], 513, and [6], 417-424). Let $J_\nu(t, \omega) = X_\nu(t, \omega) - X_\nu(t-, \omega)$. Except for a countable number of t (depending on ω), $J_\nu(t, \omega) = 0$. In addition,

$$X_\nu(t, \omega) = \sum_{\tau \leq t} J_\nu(\tau, \omega).$$

If $A \subset [0, \infty) \times (0, \infty)$ and if A is measurable, we let $N_\nu(A, \omega)$ equal the number of t for which $(t, J_\nu(t, \omega)) \in A$. Then, the random variable, $N_\nu(A, \cdot)$, has a Poisson distribution with parameter $(\lambda \times \nu)(A)$ where λ is Lebesgue measure. [If $(\lambda \times \nu)(A) = \infty$, then $P\{\omega: N_\nu(A, \omega) = \infty\} = 1$.] If $A, B \subset [0, \infty) \times (0, \infty)$, A and B are measurable and $A \cap B = \varnothing$, then $N_\nu(A, \cdot)$ and $N_\nu(B, \cdot)$ are independent random variables.

Three standard types of abbreviations will often be used: they are illustrated by the following "equalities:"

- (i) $X_\nu(t) = X_\nu(t, \omega) = X_\nu(t, \cdot)$;
- (ii) $P\{X_\nu(t) < x\} = P\{\omega: X_\nu(t, \omega) < x\}$;
- (iii) $\{\omega: X_\nu(t, \omega) \geq h(t) \text{ i.o. as } t \rightarrow 0\}$
 $= \{X_\nu(t) \geq h(t) \text{ i.o. as } t \rightarrow 0\}$
 $= \{\omega: \text{Given } T > 0, \exists t(T, \omega) \ni 0 < t(T, \omega) < T$
 $\text{and } X_\nu(t(T, \omega), \omega) \geq h(t(T, \omega))\}$.

Our final item of notation is as follows:

$$M = \{h: h \text{ is a strictly increasing function from } [0, \infty) \text{ onto } [0, \infty) \text{ and } h \text{ is concave upward}\}.$$

Note. h is concave upward means that

$$h(\lambda a + (1 - \lambda)b) \leq \lambda h(a) + (1 - \lambda)h(b), \quad 0 < \lambda < 1.$$

If $h \in M$, we write $X_{\nu \circ h}(t, \omega) = \sum_{\tau \leq t} h^{-1}(J_\nu(\tau, \omega))$.

3. Theorems and proofs. In a first reading of Theorems 1 and 2, the reader is advised to skip over statements labeled with a primed numeral.

THEOREM 1. *Let $h \in M$. Then the following statements are equivalent:*

- (i) $P\{X_\nu(t) \geq h(t) \text{ i.o. as } t \rightarrow 0\} = 0;$
 (i') $P\{X_\nu(t) > h(t) \text{ i.o. as } t \rightarrow 0\} = 0;$
- (ii) $\int_0^1 [1 - F_\nu(t, h(t))]/t dt < \infty;$
 (ii') $\int_0^1 [1 - F_\nu(t, h(t) +)]/t dt < \infty;$
- (iii) $\int_0^{1-} h^{-1}(y) \nu(dy) < \infty;$
 (iii') $\int_0^1 \nu[h(t), \infty) dt = \int_0^1 \nu(h(t), \infty) dt < \infty;$
- (iv) $\int_0^1 \left[g_\nu \left(\frac{1}{h(t)} \right) - \frac{1}{h(t)} g'_\nu \left(\frac{1}{h(t)} \right) \right] dt < \infty.$

If statements (i) and (i') are false, then true statements are obtained by replacing 0 by 1 in the right hand sides of these statements. If $h_c(t) = h(ct)$ and $H_c(t) = ch(t)$, $0 < c < \infty$, then $h_c \in M$ and $H_c \in M$, and the truth or falsity of the above statements is not changed by replacing h by h_c or H_c . The integrals in (ii) and (ii') conceivably might not exist (even in the sense of $= \infty$): if so, either upper or lower integrals may be used.

REMARK. If $h(t) = t$, then $h \in M$ and statement (iii) is true. On the other hand, if $\nu(0, \infty) = \infty$, then there exists an $h \in M$ such that statement (iii) is false. If $\nu(0, \infty) < \infty$, then $P\{X_\nu(t) > 0 \text{ i.o. as } t \rightarrow 0\} = 0$.

REMARK. Theorem 3 of [9] together with $h(t) = t^{1/2}$ shows us that statements (i) and (iii) are not equivalent if we do not assume X_ν to be increasing or $h \in M$.

REMARK. The term $\left[-\frac{1}{h(t)} g'_\nu \left(\frac{1}{h(t)} \right) \right]$ cannot be dropped from proposition (iv). One obtains a counter-example by taking $h(t) = t$ and

$$v[y, \infty) = \begin{cases} y^{-1}(\log y)^{-2}, & y \leq .1 \\ 0, & y > .1. \end{cases}$$

Proof. Part 0. We shall assume that the integrals in parts (ii) and (ii') exist (We do not exclude $= \infty$): if not, only minor changes are needed in the proof.

Part 1. Obviously (i) \Rightarrow (i') and (ii) \Rightarrow (ii').

Part 2. We prove (iii) \Leftrightarrow (iii'). Using integration by parts, we obtain

$$h^{-1}(1)\nu[1, \infty) + \int_{\delta+}^{1-} h^{-1}(y)\nu(dy) = h^{-1}(\delta)\nu(\delta, \infty) + \int_{\delta}^1 \nu(y, \infty)dh^{-1}(y) .$$

We note that

$$h^{-1}(\delta)\nu(\delta, \infty) \leq \int_0^{\delta} \nu(y, \infty)dh^{-1}(y) .$$

It follows that (iii) \Leftrightarrow (iii') once it is realized that the statement

$$\int_0^1 \nu(y, \infty)dh^{-1}(y) < \infty$$

is equivalent to (iii').

Part 3. We prove (i') \Rightarrow (iii'). Let $A = \{(t, x): t \in [0, 1], x > h(t)\}$. Then, since $X_{\nu}(t, \omega) \geq J_{\nu}(t, \omega)$, $N_{\nu}(A)$ is finite with probability one if (i') is true; and hence, if (i') is true, $(\lambda \times \nu)(A) < \infty$. But

$$(\lambda \times \nu)(A) = \int_0^1 \nu(h(t), \infty)dt .$$

Part 4. We now prove (iii) \Rightarrow (i). Since h^{-1} is concave downward, we have

$$X_{\nu \circ h}(t, \omega) = \sum_{\tau \leq t} h^{-1}(J_{\nu}(\tau, \omega)) \geq h^{-1}\left(\sum_{\tau \leq t} J_{\nu}(\tau, \omega)\right) = h^{-1}(X_{\nu}(t, \omega)) .$$

So we can prove (i) by proving

$$P\{X_{\nu \circ h}(t) \geq t \text{ i.o. as } t \rightarrow 0\} = 0 .$$

The notation $X_{\nu \circ h}$ is appropriate since it is easy to see that $X_{\nu \circ h}$ is the increasing process corresponding to the Levy measure given by

$$\mu[y, \infty) = (\nu \circ h)[y, \infty) = \nu[h(y), \infty) \quad (\text{i.e. Notation: } \mu = \nu \circ h) .$$

The only thing to be checked is that

$$\int_0^{\infty} y(1+y)^{-1}\mu(dy) < \infty .$$

But this statement is clearly equivalent to (iii).

If $P\{X_{\mu}(t) \geq t \text{ i.o. as } t \rightarrow 0\} \neq 0$, then, by Blumenthal's 0-1 law (page 57, [3]), $P\{X_{\mu}(t) \geq t \text{ i.o. as } t \rightarrow 0\} = 1$. We consider the set $\{(t_0, \omega): t_0 \in [0, 1), X_{\mu}(t_0 + t, \omega) - X_{\mu}(t_0, \omega) \geq t \text{ i.o. as } t \downarrow 0\}$. This set is measurable since it equals

$$\bigcap_{n=1}^{\infty} \bigcup_{m=n}^{\infty} \bigcap_{k=1}^{\infty} \bigcup_{r} \bigcup_{2^{-m} \leq s \leq 2^{-m+1}} \{(t_0, \omega): 0 \leq t_0, r - 2^{-k} \leq t_0 \leq r, \\ X(r+s, \omega) - X(r, \omega) \geq s - 2^{-k}\}$$

where r and s are rational. Thus, by Fubini's Theorem it follows that

$$P\{\text{Lebesgue measure of } \{t_0: t_0 \in [0, 1) \text{ and } X_\mu(t_0 + t) - X_\mu(t_0) \geq t \text{ i.o. as } t \downarrow 0\} = 1\} = 1 .$$

For each ω consider all intervals $[t_1, t_2] \subset [0, 1)$ having the property that

$$X_\mu(t_2, \omega) - X_\mu(t_1, \omega) \geq t_2 - t_1 .$$

For a subset, Ω' , of Ω having probability one, the set of all such intervals covers, in the sense of Vitali, a subset of $[0, 1)$ having Lebesgue measure equal to one. If $\varepsilon > 0$ and $\omega \in \Omega'$, we conclude, by the Vitali covering theorem that there exists a sequence

$$0 \leq t_1(\omega) < t_2(\omega) < \dots < t_{2n}(\omega) \leq 1$$

such that

$$X_\mu(t_{2k}(\omega), \omega) - X_\mu(t_{2k-1}(\omega), \omega) \geq t_{2k}(\omega) - t_{2k-1}(\omega), k = 1, \dots, n ;$$

$$\sum_{k=1}^n [t_{2k}(\omega) - t_{2k-1}(\omega)] > 1 - \varepsilon .$$

Hence, for $\omega \in \Omega'$, $X_\mu(1, \omega) \geq 1$.

Thus, we have arrived at a contradiction since $F_\mu(1, 1) > 0$ which is a consequence of the fact that X_μ has no deterministic linear component; in fact, $F_\mu(t, x) > 0$ if $x > 0$. This might be a good point to mention that it has been proved in [1] that the formulas (given at the beginning of § 2 of this paper) characterizing X_μ guarantee that X_μ has no negative deterministic linear component; but I have not seen in the literature any explicit proof of the fact that these formulas guarantee that X_μ has no positive deterministic linear component. Let us look at a proof of this fact.

If $F_\mu(t, x) = 0$ and $x > 0$, then

$$1 - \exp\left(-t \int_0^\infty (1 - e^{-uz}) \mu(dz)\right)$$

$$= \int_0^\infty (1 - e^{-uy}) d_2 F_\mu(t, y)$$

$$= \int_x^\infty (1 - e^{-uy}) d_2 F_\mu(t, y) \geq 1 - e^{-ux} .$$

Hence,

$$\int_0^\infty (1 - e^{-uz}) \mu(dz) \geq ux/t .$$

But

$$u^{-1} \int_0^\infty (1 - e^{-uz}) \mu(dz)$$

$$\leq \int_0^{(1/u)^+} z \mu(dz) + u^{-1} \mu(u^{-1}, \infty) \rightarrow 0 \text{ as } u \rightarrow \infty .$$

Thus, we have arrived at the desired contradiction.

I wish to thank Professor Steven Orey for some helpful suggestions on simplifying this part of the proof: my original proof was much more computational in nature.

Part 5. Proposition (iii') is true if and only if it is true with h replaced by h_c : this follows by a simple change of variables. Since h is concave upwards, we have

$$\begin{aligned} h(t) &\leq H_c(t) \leq h_c(t), & c \geq 1; \\ h(t) &\geq H_c(t) \geq h_c(t), & c \leq 1. \end{aligned}$$

Hence, proposition (iii') is true if and only if it is true with h replaced by H_c . Moreover, since we have now proved statements (i), (i'), and (iii) to be equivalent to (iii'), we can make similar assertions about (i), (i'), and (iii).

Part 6. We now prove (ii') \Rightarrow (i). Now (ii) implies

$$\begin{aligned} \infty &> \int_0^1 [1 - F_\nu(t, h_5(t/4))]/t \, dt = \sum_{n=0}^{\infty} \int_{2^{-n-1}}^{2^{-n}} [1 - F_\nu(t, h_5(t/4))]/t \, dt \\ &\geq \sum_{n=0}^{\infty} (2^{-n} - 2^{-n-1}) 2^n [1 - F_\nu(2^{-n-1}, h_5(2^{-n-2}))] \\ &= \frac{1}{2} \sum_{n=0}^{\infty} [1 - F_\nu(2^{-n-1}, h_5(2^{-n-2}))]. \end{aligned}$$

Hence, only finitely many of the following events can occur:

$$\{X_\nu(2^{-n-1}) \geq h_5(2^{-n-2})\}; \quad n = 0, 1, 2, \dots$$

Thus certainly only finitely many of the following events can occur:

$$\{X_\nu(t) \geq h_5(t) \text{ for some } t \in [2^{-n-2}, 2^{-n-1}]\}; \quad n = 0, 1, 2, \dots$$

Therefore, using part 5 also, we conclude that (i) is true.

Part 7. We assume (i) to be true and prove (ii). By part 5, we assume (i) to be true with h replaced by $h_{1/4}$. Hence, with probability 1, only finitely many of the following mutually independent events can occur:

$$\{X_\nu(2^{-n}) - X_\nu(2^{-n-1}) \geq h_{1/4}(2^{-n})\}, \quad n = 0, 1, 2, \dots;$$

or equivalently, the events

$$\{X_\nu(2^{-n}) - X_\nu(2^{-n-1}) \geq h(2^{-n-2})\}, \quad n = 0, 1, 2, \dots$$

Therefore,

$$\begin{aligned} \infty &> \sum_{n=0}^{\infty} [1 - F_\nu(2^{-n-1}, h(2^{-n-2}))] \\ &\geq \sum_{n=0}^{\infty} \int_{2^{-n-2}}^{2^{-n-1}} [1 - F_\nu(t, h(t))]/t \, dt = \int_0^{1/2} [1 - F_\nu(t, h(t))]/t \, dt. \end{aligned}$$

Part 8. Assume that (iv) is true. Then

$$\begin{aligned} \infty &> \int_0^1 \int_0^\infty \left[1 - \left(1 + \frac{y}{h(t)} \right) \exp \left(-\frac{y}{h(t)} \right) \right] \nu(dy) dt \\ &\cong \int_0^1 \int_{h(t)+}^\infty \left[1 - \left(1 + \frac{y}{h(t)} \right) \exp \left(-\frac{y}{h(t)} \right) \right] \nu(dy) dt \\ &\cong (1 - 2e^{-1}) \int_0^1 \nu(h(t), \infty) dt . \end{aligned}$$

At the last step we used the fact that $(1 + a)e^{-a}$ decreases as a increases. We have proved that (iv) \Rightarrow (iii').

Part 9. We assume that (iii) and (iii') are true and prove (iv). We note that $1 - (1 + a)e^{-a} \leq a^2/2$ if $a \geq 0$. Then,

$$\begin{aligned} &\int_0^1 \left[g_\nu \left(\frac{1}{h(t)} \right) - \frac{1}{h(t)} g'_\nu \left(\frac{1}{h(t)} \right) \right] dt \\ &= \int_0^1 \int_0^\infty \left[1 - \left(1 + \frac{y}{h(t)} \right) \exp \left(-\frac{y}{h(t)} \right) \right] \nu(dy) dt \\ &= \int_0^1 \int_0^{h(t)+} \left[1 - \left(1 + \frac{y}{h(t)} \right) \exp \left(-\frac{y}{h(t)} \right) \right] \nu(dy) dt \\ &\quad + \int_0^1 \int_{h(t)+}^\infty \left[1 - \left(1 + \frac{y}{h(t)} \right) \exp \left(-\frac{y}{h(t)} \right) \right] \nu(dy) dt \\ &\cong \int_0^1 \int_0^{h(t)+} \frac{y^2}{2[h(t)]^2} \nu(dy) dt + \int_0^1 \nu(h(t), \infty) dt . \end{aligned}$$

The second term is finite. Since h is concave upward, the first term is no larger than

$$\begin{aligned} &\frac{1}{2} \int_0^1 \int_0^{h(t)+} \frac{[h^{-1}(y)]^2}{t^2} \nu(dy) dt \\ &= \frac{1}{2} \int_0^{h(1)+} \int_{h^{-1}(y)}^1 \frac{[h^{-1}(y)]^2}{t^2} dt \nu(dy) \\ &= \frac{1}{2} \int_0^{h(1)+} h^{-1}(y) [1 - h^{-1}(y)] \nu(dy) < \infty . \end{aligned}$$

Part 10. Part 3 of the proof shows that if (iii') is false, then a true statement is obtained by replacing 0 by 1 in the right hand side of (i'); and, thus, in the right hand side of (i). One could also use Blumenthal's 0-1 law on page 57 of [3] to arrive at this conclusion. The proof is complete.

If one is "given" an increasing process with stationary independent increments—i.e. if one is given ν , F_ν , g_ν , or φ_ν —and if one is given $h \in M$, one might quite easily ascertain whether or not one of the statements (ii), (iii), or (iv) is true; thus one might easily conclude

whether or not statements (i) and (i') are true. We now prove a similar theorem concerning the behavior of the sample paths for large values of t . The statement of the theorem is somewhat complicated by the fact that the behavior for large t can depend on both the "small" and "large" jumps; whereas the behavior for small t depends only on the "small" jumps.

THEOREM 2. *Let $h \in M$. Let $\alpha \in (0, \infty)$ be defined by the equation $\alpha = \inf \{c: ct \geq h(t) \text{ for } t \in [0, \infty)\}$. If $\alpha < \infty$, then the following statements are equivalent:*

- (i) $P\{X_\nu(t) \geq h(t) \text{ i.o. as } t \rightarrow \infty\} = 0$;
- (i') $P\{X_\nu(t) > h(t) \text{ i.o. as } t \rightarrow \infty\} = 0$;
- (aii) $\int_0^\infty x d_2 F_\nu(t, x) < \alpha t$ for one (all) $t \in (0, \infty)$;
- (aiii) $\int_0^\infty y \nu(dy) = \int_0^\infty \nu[y, \infty) dy = \int_0^\infty \nu(y, \infty) dy < \alpha$;
- (aiv) $g'_\nu(0) < \alpha$.

If $\alpha = \infty$, then the following statements are equivalent:

- (i) $P\{X_\nu(t) \geq h(t) \text{ i.o. as } t \rightarrow \infty\} = 0$;
- (i') $P\{X_\nu(t) > h(t) \text{ i.o. as } t \rightarrow \infty\} = 0$;
- (bii) $\int_1^\infty [1 - F_\nu(t, h(t))]/t dt < \infty$;
- (bii') $\int_1^\infty [1 - F_\nu(t, h(t) +)]/t dt < \infty$;
- (biii) $\int_{1-}^\infty h^{-1}(y) \nu(dy) < \infty$;
- (biii') $\int_1^\infty \nu[h(t), \infty) dt = \int_1^\infty \nu(h(t), \infty) dt < \infty$;
- (biv) $\int_1^\infty \left[g_\nu\left(\frac{1}{h(t)}\right) - \frac{1}{h(t)} g'_\nu\left(\frac{1}{h(t)}\right) \right] dt < \infty$.

If statements (i) and (i') are false, then true statements are obtained by replacing 0 by 1 in the right hand sides of these statements. If $\alpha = \infty$, then the truth or falsity of (i), (i'), (bii), (bii'), (biii), (biii'), and (biv) is not changed by replacing h by h_c or H_c . The integrals in (bii) and (bii') conceivably might not exist (even in the sense of $= \infty$): if so, either upper or lower integrals may be used.

REMARK. There exists $h \in M$ such that (i) is false. Also, there exists $h \in M$ such that (i) is true.

Proof. Part 0. We shall assume that the integrals in (bii) and (bii') exist (We do not exclude $= \infty$): if not, only minor changes are needed in the proof.

Part 1. Obviously (i) \Rightarrow (i') and (bii) \Rightarrow (bii').

Part 2. Note that $\int_0^{1-} y\nu(dy) < \infty$. Using this fact together with integration by parts (analogous to part 2 of the proof of Theorem 1), we easily deduce the two equalities in (aiii) and the equivalence of (biii) and (biii').

Part 3. We complete the proof in case $\alpha < \infty$. We have

$$X_\nu(nt, \omega) = \sum_{k=1}^n X_\nu(kt, \omega) - X_\nu((k-1)t, \omega).$$

The random variables $X_\nu(kt) - X_\nu((k-1)t), k = 1, \dots$ are mutually independent and identically distributed with distribution function $F_\nu(t, \cdot)$. Hence, by the law of large numbers,

$$P\left\{\lim_{n \rightarrow \infty} \frac{X_\nu(nt)}{n} = \int_0^\infty x d_2 F_\nu(t, x)\right\} = 1.$$

By Theorem 4 of [5] we can, in fact, say that if

$$\int_0^\infty x d_2 F_\nu(t, x) < \infty,$$

then

$$P\left\{X_\nu(nt) > n \int_0^\infty x d_2 F_\nu(t, x) \text{ i.o. as } n \rightarrow \infty\right\} = 1.$$

Hence, using the fact that $h \in M$, the truth of the theorem for $\alpha < \infty$ will follow once it is shown that

$$t \int_0^\infty y \nu(dy) = t g'_\nu(0) = \int_0^\infty x d_2 F_\nu(t, x) \text{ (possibly } + \infty \text{)}.$$

We have

$$e^{-t g_\nu(u)} = \int_0^\infty e^{-ux} d_2 F_\nu(t, x).$$

For $u > 0$, we can differentiate under the integral to obtain

$$-t g'_\nu(u) e^{-t g_\nu(u)} = - \int_0^\infty x e^{-ux} d_2 F_\nu(t, x). \text{ Now let } u \rightarrow 0$$

and use the fact that g_ν is monotone. Hence, the right hand equality follows. We also have $g_\nu(u) = \int_0^\infty (1 - e^{-uy}) \nu(dy)$. Differentiating and letting $y \rightarrow 0$, we obtain the left hand equality.

Part 4. We imitate part 3 of the proof Theorem 1 to show that (i') \Rightarrow (biii') and that, if (biii') is false a true statement is obtained by replacing 0 by 1 in the right hand side of (i'). Theorem 11.3 of [8], although not necessary, could be used here.

Part 5. We assume that (biii') is true and $\alpha = \infty$ and we prove (i). Let $\beta > 0$ be such that $\int_{\beta-}^{\infty} h^{-1}(y)\nu(dy) < 1/2$. We define measures η and ξ on $(0, \infty)$ by the formulae

$$\begin{aligned}\eta(B) &= \nu(B \cap (0, \beta)) ; \\ \xi(B) &= \nu(B \cap [\beta, \infty)) .\end{aligned}$$

By what we proved in part 3 it follows that

$$\begin{aligned}P\{X_{\eta}(t) \geq h(t)/2 \text{ i.o. as } t \rightarrow \infty\} \\ \leq P\left\{X_{\eta}(t) \geq 2t \int_0^{\beta-} y\eta(dy) \text{ i.o. as } t \rightarrow \infty\right\} = 0 .\end{aligned}$$

Hence, it will suffice to show that

$$P\{X_{\xi}(t) \geq h(t)/2 \text{ i.o. as } t \rightarrow \infty\} = 0 .$$

But this will follow if we can show that

$$P\{X_{\xi \circ (h/2)}(t) \geq t \text{ i.o. as } t \rightarrow \infty\} = 0 ;$$

for, as in part 4 of the proof of Theorem 1, we have

$$X_{\xi \circ (h/2)}(t, \omega) \geq h^{-1}(2X_{\xi}(t, \omega)) .$$

We note that

$$\begin{aligned}\int_0^{\infty} y(\xi \circ (h/2))(dy) &= \int_0^{\infty} h^{-1}(2y)\xi(dy) \\ &\leq 2 \int_0^{\infty} h^{-1}(y)\xi(dy) = 2 \int_{\beta-}^{\infty} h^{-1}(y)\nu(dy) < 1 .\end{aligned}$$

The desired result follows by what was proved in part 3.

The author would like to thank the referee for pointing out an error in the original calculations above. Also, he made several helpful suggestions about the organization of the paper.

Part 6. As in part 5 of the proof of Theorem 1, we can show that, if $\alpha = \infty$, then the truth or falsity of (i), (i'), (biii), and (biii') is unchanged if h is replaced by h_e or H_e .

Part 7. That (bii') \Rightarrow (i) \Rightarrow (bii) when $\alpha = \infty$ can easily be shown using methods analogous to those used in parts 6 and 7 of the proof of Theorem 1.

Part 8. Let $\alpha = \infty$. Then to show that (biii) \Leftrightarrow (biv) we proceed in a manner similar to parts 8 and 9 of the proof of Theorem 1. At one point the calculations become dissimilar: the appropriate calculation follows:

$$\begin{aligned} & \frac{1}{2} \int_1^\infty \int_0^{h(t)+} \frac{y^2}{[h(t)]^2} \nu(dy) dt \\ &= \frac{1}{2} \left[\int_0^{h(1)+} \int_1^\infty + \int_{h(1)+}^\infty \int_{h^{-1}(y)}^\infty \right] \frac{y^2}{[h(t)]^2} dt \nu(dy) . \end{aligned}$$

Now

$$\left[\int_0^{h(1)+} y^2 \nu(dy) \right] \left[\int_1^\infty \frac{1}{[h(t)]^2} dt \right] < \infty .$$

Since $h \in M$, we have

$$\begin{aligned} & \int_{h(1)+}^\infty \int_{h^{-1}(y)}^\infty \frac{y^2}{[h(t)]^2} dt \nu(dy) \leq \int_{h(1)+}^\infty \int_{h^{-1}(y)}^\infty \frac{[h^{-1}(y)]^2}{t^2} dt \nu(dy) \\ &= \int_{h(1)+}^\infty h^{-1}(y) \nu(dy) < \infty . \end{aligned}$$

The proof of Theorem 2 is complete.

We now prove a variation-type theorem.

THEOREM 3. *Let $h^{-1} \in M$ and $t \in (0, \infty)$. For almost all $\omega \in \Omega$ the following statements are true. If $\int_0^1 h(y) \nu(dy) = \infty$, then, given $L < +\infty$, there exists $\delta(= \delta(t, \omega)) > 0$ such that*

$$\sum_{j=1}^n h(X_\nu(t_j) - X_\nu(t_{j-1})) > L$$

whenever $0 \leq t_j - t_{j-1} < \delta$ and $t_0 = 0, t_n = t$. If $\int_0^1 h(y) \nu(dy) < \infty$, then, given $\varepsilon > 0$, there exists $\delta(= \delta(t, \omega)) > 0$ such that

$$0 \leq X_{\nu \circ h^{-1}}(t) - \sum_{j=1}^n h(X_\nu(t_j) - X_\nu(t_{j-1})) < \varepsilon$$

whenever $0 \leq t_j - t_{j-1} < \delta$ and $t_0 = 0, t_n = t$.

Proof. Let $\mu = \nu \circ h^{-1}$. Because h is concave downward it follows that $X_\mu(t) - \sum_{j=1}^n h(X_\nu(t_j) - X_\nu(t_{j-1})) \geq 0$. If $\varepsilon > 0$ is given then we can find $s_1, \dots, s_m < t$ such that $X_\mu(t) - \sum_{i=1}^m h(J_\nu(s_i)) < \varepsilon$. We certainly can find δ [even if $\int_0^1 h(y) \nu(dy) = \infty$] such that

$$\sum_{j=1}^n h(X_\nu(t_j) - X_\nu(t_{j-1})) \geq \sum_{i=1}^m h(J_\nu(s_i))$$

whenever $0 \leq t_j - t_{j-1} < \delta$ and $t_0 = 0, t_n = t$. If $\int_0^1 h(y) \nu(dy) = \infty$, we might still define $X_\mu(t) = \sum_{\tau \leq t} h(J_\nu(\tau))$, although $X_\mu(t)$ might conceivably equal $+\infty$. If $X_\mu(t) = +\infty$, then we can find $s_1, \dots, s_m < t$ such that $\sum_{i=1}^m h(J_\nu(s_i)) > L$. Putting the statements of the above paragraph

together we see that we are finished once we prove that $P\{X_\mu(t) = \infty\} = 1$ assuming that $\int_0^1 h(y)\nu(dy) = \infty$. We let $N_j(t, \omega)$ equal the number of $\tau \in [0, t]$ having the property that $J_\nu(\tau, \omega) \in [2^{-j-1}, 2^{-j}]$; $j = 0, 1, \dots$. Then

$$X_\mu(t) \geq \sum_{j=0}^{\infty} h(2^{-j-1})N_j(t) \geq \frac{1}{2} \sum_{j=0}^{\infty} h(2^{-j})N_j(t).$$

Hence,

$$\begin{aligned} E\left\{\exp\left[-2X_\mu(t)\right]\right\} &\leq E\left\{\exp\left[-\sum_{j=0}^{\infty} h(2^{-j})N_j(t)\right]\right\} \\ &= \prod_{j=0}^{\infty} E\{\exp[-h(2^{-j})N_j(t)]\} \\ &= \prod_{j=0}^{\infty} \exp\{-t\nu[2^{-j-1}, 2^{-j}][1 - \exp(-h(2^{-j}))]\} \\ &= \exp\left\{-t \sum_{j=0}^{\infty} [1 - \exp(-h(2^{-j}))]\nu[2^{-j-1}, 2^{-j}]\right\} \\ &\leq \exp\left\{-t \int_0^1 [1 - e^{-h(y)}]\nu(dy)\right\} \\ &\leq \exp\left\{-t (\text{const.}) \int_0^1 h(y)\nu(dy)\right\} = 0; \end{aligned}$$

where we have used the fact that if U is Poisson distributed with parameter η and if a is some nonnegative constant, then

$$E\{e^{-aU}\} = \exp[-\eta(1 - e^{-a})].$$

We now prove an L_1 convergence theorem.

THEOREM 4. *Let $h^{-1} \in M$ and assume that $\int_0^1 \nu[y, \infty)dh(y) < \infty$. Then*

$$\lim_{t \rightarrow 0} \int_0^y \left| \frac{1 - F_\nu(t, x)}{t} - \nu[x, \infty) \right| dh(x) = 0.$$

REMARK. By the central convergence criterion ([11], p. 311), we know that $\lim_{t \rightarrow 0} [1 - F_\nu(t, y)]/t = \nu[y, \infty)$ for every continuity point y of $\nu[\cdot, \infty)$.

Proof. Let $G(t, y) = F_\nu(t, h^{-1}(y))$ and $\mu[y, \infty) = \nu[h^{-1}(y), \infty)$. Then Theorem 3 together with the central convergence criterion ([11], p. 311) gives us

$$\begin{aligned} \lim_{t \rightarrow 0} t^{-1} \int_0^{y^-} h(x) d_2 F_\nu(t, x) &= \lim_{t \rightarrow 0} t^{-1} \int_0^{h(y)^-} x d_2 G(t, x) \\ &= \int_0^{h(y)} x \mu(dx) = \int_0^y h(x) \nu(dx) \end{aligned}$$

for any y which is a continuity point of $\nu[\cdot, \infty)$. Integration by parts together with the remark preceding this proof gives us

$$\lim_{t \rightarrow 0} t^{-1} \int_0^y [1 - F_\nu(t, x)] dh(x) = \int_0^y \nu[x, \infty) dh(x).$$

The remark after the theorem together with the fact just proved implies the desired conclusion.

Except for the assertions involving $g(u)$, the four theorems generalize in the obvious way to the situation where ν is a measure on R^N such that $\int_{0 < |x| < 1} |x| \nu(dx) < \infty$.

BIBLIOGRAPHY

1. Glen Baxter and J. M. Shapiro, *On bounded infinitely divisible random variables*, Sankhyā: The Indian Journal of Statistics, **22** (1960), 253-260.
2. D. Bierens de Hahn, *Nouvelles Tables d'Intégrales Définies*, G. E. Stechert and Company, 1939 (corrected the edition of 1867).
3. R. M. Blumenthal, *An extended Markov property*, Trans. Amer. Math. Soc. **85** (1957), 52-72.
4. R. M. Blumenthal and R. K. Gettoor, *Sample function of stochastic processes with independent increments*, J. Math. Mech. **10** (1961), 493-516.
5. K. L. Chung and W. H. J. Fuchs, *On the distribution of values of sums of random variables*, Number 6, Memoirs of the Amer. Math. Soc., 1951.
6. J. L. Doob, *Stochastic Processes*, John Wiley and Sons, Inc., London, 1953.
7. Bert Fristedt, *Generalized variations, continuity, and Hausdorff measures of sample functions of stochastic processes with independent increments*, Thesis, MIT, 1963.
8. Edwin Hewitt and Leonard J. Savage, *Symmetric measures on cartesian products*, Trans. Amer. Math. Soc. **80** (1955), 470-501.
9. A. Khinchin, *Sur la croissance locale des processus stochastiques homogènes à accroissements indépendants* (Russian article and French resume), Akad. Nauk. SSSR, Izvestia, Ser. Math. **3** (1939), 487-508.
10. ———, *Zwei Sätze über stochastische Prozesse mit stabilen Verteilungen* (Russian article and German resume), Matematicheskii Sbornik **3** (1938), New Series, 577-584.
11. Michel Loève, *Probability Theory*, 2nd edition, D. Van Nostrand Company, Inc., Princeton, New Jersey, 1960.

Received November 18, 1965. This research was supported in part by the Air Force Office of Scientific Research under grant Af-AFOSR 62-252. Theorem 3 together with some of the easier parts of the proof of Theorem 1 are contained in the author's Ph. D. thesis completed at MIT under the direction of Henry P. McKean, Jr.

CARLETON COLLEGE, NORTHFIELD, MINNESOTA
AND
UNIVERSITY OF MINNESOTA

PACIFIC JOURNAL OF MATHEMATICS

EDITORS

H. SAMELSON
Stanford University
Stanford, California

J. P. JANS
University of Washington
Seattle, Washington 98105

J. DUGUNDJI
University of Southern California
Los Angeles, California 90007

RICHARD ARENS
University of California
Los Angeles, California 90024

ASSOCIATE EDITORS

E. F. BECKENBACH

B. H. NEUMANN

F WOLF

K. Yosida

SUPPORTING INSTITUTIONS

UNIVERSITY OF BRITISH COLUMBIA
CALIFORNIA INSTITUTE OF TECHNOLOGY
UNIVERSITY OF CALIFORNIA
MONTANA STATE UNIVERSITY
UNIVERSITY OF NEVADA
NEW MEXICO STATE UNIVERSITY
OREGON STATE UNIVERSITY
UNIVERSITY OF OREGON
OSAKA UNIVERSITY
UNIVERSITY OF SOUTHERN CALIFORNIA

STANFORD UNIVERSITY
UNIVERSITY OF TOKYO
UNIVERSITY OF UTAH
WASHINGTON STATE UNIVERSITY
UNIVERSITY OF WASHINGTON
* * *
AMERICAN MATHEMATICAL SOCIETY
CHEVRON RESEARCH CORPORATION
TRW SYSTEMS
NAVAL ORDNANCE TEST STATION

Pacific Journal of Mathematics

Vol. 21, No. 1

November, 1967

Friedrich-Wilhelm Bauer, <i>Der Hurewicz-Satz</i>	1
D. W. Dubois, <i>A note on David Harrison's theory of preprimes</i>	15
Bert E. Fristedt, <i>Sample function behavior of increasing processes with stationary, independent increments</i>	21
Minoru Hasegawa, <i>On the convergence of resolvents of operators</i>	35
Søren Glud Johansen, <i>The descriptive approach to the derivative of a set function with respect to a σ-lattice</i>	49
John Frank Charles Kingman, <i>Completely random measures</i>	59
Tilla Weinstein, <i>Surfaces harmonically immersed in E^3</i>	79
Hikosaburo Komatsu, <i>Fractional powers of operators. II. Interpolation spaces</i>	89
Edward Milton Landesman, <i>Hilbert-space methods in elliptic partial differential equations</i>	113
O. Carruth McGehee, <i>Certain isomorphisms between quotients of a group algebra</i>	133
DeWayne Stanley Nymann, <i>Dedekind groups</i>	153
Sidney Charles Port, <i>Hitting times for transient stable processes</i>	161
Ralph Tyrrell Rockafellar, <i>Duality and stability in extremum problems involving convex functions</i>	167
Philip C. Tonne, <i>Power-series and Hausdorff matrices</i>	189