

Finiteness of outer automorphism groups of random right-angled Artin groups

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We consider the outer automorphism group $\text{Out}(A_\Gamma)$ of the right-angled Artin group A_Γ of a random graph Γ on n vertices in the Erdős–Rényi model. We show that the functions $n^{-1}(\log(n) + \log(\log(n)))$ and $1 - n^{-1}(\log(n) + \log(\log(n)))$ bound the range of edge probability functions for which $\text{Out}(A_\Gamma)$ is finite: if the probability of an edge in Γ is strictly between these functions as n grows, then asymptotically $\text{Out}(A_\Gamma)$ is almost surely finite, and if the edge probability is strictly outside of both of these functions, then asymptotically $\text{Out}(A_\Gamma)$ is almost surely infinite. This sharpens a result of Ruth Charney and Michael Farber.

[05C80](#), [20E36](#), [20F28](#), [20F69](#); [20F05](#)

1 Introduction

Let Γ be a simplicial graph with vertex set V . The *right-angled Artin group* A_Γ defined by Γ is the group with the presentation

$$\langle V \mid ab = ba \text{ for all } a, b \in V \text{ with } a \text{ adjacent to } b \rangle.$$

Right-angled Artin groups include free groups and free abelian groups and are common objects of study in geometric group theory. Outer automorphism groups of right-angled Artin groups exhibit great variety: although they include infinite groups such as outer automorphism groups of free groups and $\text{GL}(n, \mathbb{Z})$, many of them are finite.

The theory of random graphs is a branch of combinatorics initiated by Erdős and Rényi in a 1959 paper [4]. Since right-angled Artin groups are indexed over graphs, it is natural to ask about the properties of random ones. Random right-angled Artin groups were studied by Costa–Farber in [3], and their automorphism groups were specifically studied by Charney–Farber in [2]. Charney and Farber showed that under certain conditions, a random right-angled Artin group almost certainly has a finite outer automorphism group; the results of this paper are a sharpening of their results.

1.1 Background

This paper is about two graph-theoretic notions that arise in the study of right-angled Artin groups: *domination* and *star 2-connectedness*. Again let Γ be a simplicial graph with vertex set V and adjacency relation \sim . The *star* of a vertex a of Γ is the set consisting of a and all vertices adjacent to a :

$$\text{st}(a) = \{a\} \cup \{b \in V \mid b \sim a\}.$$

A vertex $a \in V$ is a *star-cut-vertex* for Γ if the full subgraph $\Gamma \setminus \text{st}(a)$ is disconnected. The graph Γ is *star 2-connected* if it has no star-cut-vertices. If Γ is star 2-connected, then either Γ is connected or Γ is the disjoint union of exactly two complete graphs. For a pair of distinct vertices $a, b \in V$, a *dominates* b in Γ if every vertex adjacent to b is adjacent to or equal to a , in other words that

$$\text{st}(b) \setminus \{b\} \subset \text{st}(a).$$

We write $a > b$ if a dominates b , and refer to (a, b) as a *domination pair*. Note that it is possible for a to dominate b whether $a \sim b$ or not; if $a \sim b$ then (a, b) is an *adjacent domination pair*, and otherwise it is a *nonadjacent domination pair*. A vertex is *isolated* if it is adjacent to no other vertices, and it is *central* if it is adjacent to all other vertices. Note that if a is isolated then $b > a$ for any b , and if a is central, then $a > b$ for any b . The following examples are instructive: a path graph (that is, an unbranching tree) with at least four vertices has exactly four domination pairs, and a path graph with at least five vertices is not star 2-connected, but a cycle graph on at least five vertices is star 2-connected and has no domination pairs.

The presence of domination pairs and star-cut-vertices indicate the existence of infinite order outer automorphisms of right-angled Artin groups. Laurence showed in [5] that the automorphism group $\text{Aut}(A_\Gamma)$ of the right-angled Artin group A_Γ of a finite graph Γ is generated by finitely many automorphisms that fall into four classes: inversions, symmetries, dominated transvections, and partial conjugations. While inversions and symmetries generate a finite subgroup of $\text{Aut}(A_\Gamma)$, dominated transvections and partial conjugations are infinite order. A dominated transvection always has an infinite-order image in the outer automorphism group $\text{Out}(A_\Gamma)$, but a dominated transvection will only exist if Γ has a domination pair. If Γ is star 2-connected, then every partial conjugation is an inner automorphism; if there is a star-cut-vertex, then there is a partial conjugation whose image in $\text{Out}(A_\Gamma)$ has infinite order. We have explained the following fact.

Fact 1.1 *The group $\text{Out}(A_\Gamma)$ is finite if and only if Γ is star 2-connected and has no domination pairs.*

For the details of this argument, we refer the reader to Charney–Farber [2, Section 6].

To proceed, we must formalize our notion of random graphs. The *Erdős–Rényi model* for random graphs is the sequence of probability spaces $G(n, p)$, where

- n varies over the positive integers,
- $p = p(n)$ is a sequence of probability values in $[0, 1]$,
- the underlying set of $G(n, p)$ is the finite set of all simplicial graphs with vertex set V of cardinality $|V| = n$,
- each edge occurs with probability p and independently of other edges.

It is easy to see that this last condition uniquely determines $G(n, p)$; it means that $G(n, p)$ assigns each graph Γ with m edges the probability

$$\text{Prob}(\Gamma) = p^m(1 - p)^{\binom{n}{2} - m}.$$

Suppose we are given a sequence of probabilities p and a property of graphs P . We say that $\Gamma \in G(n, p)$ has the property P *asymptotically almost surely (a.a.s.)* if the probability that $\Gamma \in G(n, p)$ has P goes to 1 as $n \rightarrow \infty$. This model for random graphs and related models are described in detail in Bollobás [1, Chapter 2].

The work in this paper is inspired by the following results of Charney–Farber [2]:

Theorem 1.2 (Charney–Farber) *Suppose p is any function satisfying*

$$p(1 - p)n - 2 \log(n) \rightarrow \infty \quad \text{as } n \rightarrow \infty,$$

(for example, if p is constant in n with $0 < p < 1$). Then $\Gamma \in G(n, p)$ a.a.s. has no domination pairs.

Theorem 1.3 (Charney–Farber) *Suppose p is constant with respect to n and*

$$1 - \frac{1}{\sqrt{2}} < p < 1.$$

Then $\Gamma \in G(n, p)$ a.a.s. is star 2–connected.

In this paper, we find sharper descriptions of the functions p for which $\Gamma \in G(n, p)$ a.a.s. has no domination pairs and is star 2–connected. Further, we show that for p outside of these ranges, the negations of these statements hold a.a.s.

1.2 Statement of results

Our two main theorems explain the asymptotically almost sure existence and nonexistence of domination pairs and star-cut-vertices.

Theorem 1.4 *Let $C > 0$ be any fixed constant, and suppose $p = p(n)$ is a sequence of probability values. The existence of domination pairs in $\Gamma \in G(n, p)$, asymptotically almost surely, is summarized as follows:*

- *If $pn^2 \rightarrow 0$, then there are no adjacent domination pairs.*

- *If*

$$p < \frac{\log(n) + \log(\log(n)) - \omega(n)}{n}$$

for some sequence $\omega(n)$ with $\omega(n) \rightarrow +\infty$, then there are at least C nonadjacent domination pairs.

- *If we still have $p < n^{-1}(\log(n) + \log(\log(n)) - \omega(n))$, but in addition we have $pn^2 \rightarrow \infty$, then there are also at least C adjacent domination pairs.*

- *If*

$$\frac{\log(n) + \log(\log(n)) + \omega_1(n)}{n} < p < 1 - \frac{\log(n) + \log(\log(n)) + \omega_2(n)}{n},$$

for some sequences $\omega_1(n), \omega_2(n)$, both tending to positive infinity, then there are no domination pairs.

- *If*

$$p > 1 - \frac{\log(n) + \log(\log(n)) - \omega(n)}{n}$$

for some sequence $\omega(n)$ with $\omega(n) \rightarrow +\infty$, then there are at least C adjacent domination pairs.

- *If we still have $p > 1 - n^{-1} \log(n) + \log(\log(n)) - \omega(n)$, but in addition we have $(1 - p)n^2 \rightarrow \infty$, then there are also at least C nonadjacent domination pairs.*

- *If $(1 - p)n^2 \rightarrow 0$, then there are no nonadjacent domination pairs.*

Proof It is well known that if $pn^2 \rightarrow 0$, then the probability that $\Gamma \in G(n, p)$ is the edgeless graph goes to 1 (since Γ has $\binom{n}{2}$ pairs of vertices, the probability that Γ is the edgeless graph is $(1 - p)^{n(n-1)/2} \sim e^{-pn(n-1)/2}$ as $n \rightarrow \infty$). Adjacent domination pairs require the existence of edges, and therefore we have proven the first item. The last item follows by a dual argument: if $(1 - p)n^2 \rightarrow 0$, then the probability that Γ is the complete graph goes to 1, but nonadjacent domination pairs require pairs of vertices with no edges between them.

We prove the items that assert existence in [Proposition 2.7](#) and [Theorem 2.13](#) covers the item asserting the nonexistence of domination pairs. \square

Theorem 1.5 *Suppose $p = p(n)$ is a sequence of probability values. There are a.a.s. no star-cut-vertices in $\Gamma \in G(n, p)$ if*

- for some sequence $\omega(n)$ with $\omega(n) \rightarrow \infty$,

$$p > \frac{\log(n) + \log(\log(n)) + \omega(n)}{n},$$

- either $n(1 - p) \rightarrow 0$ or $n(1 - p) \rightarrow \infty$.

Further, if only the first hypothesis holds, then a.a.s. for any star-cut-vertex $a \in \Gamma$, there is at most one component of $\Gamma \setminus \text{st}(a)$ with more than one vertex.

The proof of this theorem appears in [Section 3.3](#) below. If vertices $a, b, c \in V$ form an isolated triangle in the complement graph $\bar{\Gamma}$, then each of a, b and c is a star-cut-vertex. Isolated triangles are only asymptotically forbidden if $np \rightarrow \infty$ or $np \rightarrow 0$. This fact is explained in Bollobás [[1](#), Theorem V.16, page 111]. In particular, the presence of isolated triangles in $\bar{\Gamma}$ explains the possibility of star-cut-vertices if the second hypothesis in [Theorem 1.5](#) fails.

The following corollary is the goal of the paper.

Corollary 1.6 *If the probability sequence p satisfies*

$$\frac{\log(n) + \log(\log(n)) + \omega_1(n)}{n} < p < 1 - \frac{\log(n) + \log(\log(n)) + \omega_2(n)}{n}$$

for some sequences $\omega_1, \omega_2 \rightarrow +\infty$, then $\text{Out}(A_\Gamma)$ is a.a.s. finite for $\Gamma \in G(n, p)$.

Conversely, if

$$p < \frac{\log(n) + \log(\log(n)) - \omega(n)}{n} \quad \text{or} \quad p > 1 - \frac{\log(n) + \log(\log(n)) - \omega(n)}{n}$$

for some $\omega \rightarrow +\infty$, then $\text{Out}(A_\Gamma)$ is a.a.s. infinite for $\Gamma \in G(n, p)$.

Proof [Theorem 1.4](#) explains that there are a.a.s. no domination pairs in the first case, and a.a.s. there exist some domination pairs in the second case. [Theorem 1.5](#) implies that a.a.s. there are no star-cut-vertices in the first case above. Then the corollary follows from [Fact 1.1](#). \square

Remark If Γ has isolated vertices then $\text{Out}(A_\Gamma)$ has a subgroup isomorphic to the automorphism group of a free group, and if Γ has central vertices then $\text{Out}(A_\Gamma)$ has a subgroup isomorphic to a general linear group over the integers. By a famous theorem of Erdős and Rényi (see [Theorem 2.6](#)), Γ will a.a.s. have isolated vertices if p is less than $n^{-1}(\log(n) - \omega(n))$ and central vertices if p is greater than $1 - n^{-1}(\log(n) - \omega(n))$ for some $\omega \rightarrow \infty$. However, there are two narrow ranges of probability functions, where p or $1 - p$ is between $n^{-1}(\log(n) + \omega_1(n))$ and $n^{-1}(\log(n) + \log(\log(n)) - \omega_2(n))$ for any $\omega_1, \omega_2 \rightarrow \infty$, such that Γ and $\bar{\Gamma}$ are a.a.s. connected but $\text{Out}(A_\Gamma)$ is a.a.s. infinite.

We end this section with a corollary that gives some insight into the group theory of $\text{Out}(A_\Gamma)$ in the case that p does not go to zero quickly enough.

Corollary 1.7 *If $p > n^{-1}(\log(n) + \log(\log(n)) + \omega(n))$ for some $\omega \rightarrow \infty$, then a.a.s. $\text{Out}(A_\Gamma)$ is generated by dominated transvections, symmetries, and inversions only—partial conjugations are unnecessary.*

Proof It is enough to explain why a partial conjugation can be expressed as a product of dominated transvections under this hypothesis. As explained in [Theorem 1.5](#), this hypothesis implies that a.a.s. the star of any star-cut-vertex in Γ has at most one complementary component with more than one vertex. Suppose $a \in \Gamma$ is a star-cut-vertex with this property, and $S \subset \Gamma$ is a union of complementary components of $\text{st}(a)$. The data of a and S determine a partial conjugation automorphism $\alpha \in \text{Aut}(A_\Gamma)$, which is defined on generators of A_Γ by

$$\alpha(b) = \begin{cases} aba^{-1} & b \in S, \\ b & b \notin S. \end{cases}$$

If $b \in \Gamma$ is an isolated vertex in $\Gamma \setminus \text{st}(a)$, then a dominates b and the dominated transvections multiplying b by $a^{\pm 1}$ on the right and on the left all exist; these are the four automorphisms that fix all generators other than b , but send b to $a^{\pm 1}b$ or $ba^{\pm 1}$. Let S_1 be the component of $\Gamma \setminus \text{st}(a)$ with more than one vertex, if it exists. If α does not fix S_1 (meaning that $S_1 \subset S$), we can compose α with an inner automorphism to get an automorphism that does fix S_1 . In any case, the class of α in $\text{Out}(A_\Gamma)$ is represented by an automorphism that conjugates certain generators by $a^{\pm 1}$ and fixes the rest, and such that those generators that it does not fix are all dominated by a . This automorphism is certainly a product of the dominated transvections given above. \square

1.3 General probability background

This paper is written by a group theorist, for group theorists, even though the paper consists mostly of arguments in probabilistic combinatorics. Although hopefully the

content is of some interest to combinatorialists, the arguments are intended to be intelligible to people with limited background in probability or combinatorics. The main reference on combinatorics used in preparation of the paper is Bollobás' *Random Graphs* [1].

Markov's inequality is a simple probabilistic technique used repeatedly in this paper. If X is a random variable (that is, a function to \mathbb{R}) on a probability space and $\lambda > 0$ is a real number, then we have the following bound on the expectation $\mathbb{E}(X)$ of X :

$$\lambda \text{Prob}(X \geq \lambda) \leq \mathbb{E}(X).$$

This is often expressed as $\text{Prob}(X \geq \lambda) \leq \mathbb{E}(X)/\lambda$. This appears in Bollobás [1, page 2].

The usual procedure to show asymptotic nonexistence results for some feature of graphs is as follows. We find a sequence of integer-valued random variables $X = X_n$ on the spaces $G(n, p)$, defined the same way for all n , where the desired property is true of a graph $\Gamma \in G(n, p)$ if $X(\Gamma) = 0$. Fixing n , we express X as a sum of Bernoulli random variables, that is, a sum of characteristic functions. The expectation of a Bernoulli random variable is the probability of the set on which it is nonzero, and we use linearity of expectations to find or bound the expectation $\mathbb{E}(X)$ as a function of n . Then we take the limit $\lim_{n \rightarrow \infty} \mathbb{E}(X)$ and show that it is 0, and use Markov's inequality (with $\lambda = 1/2$, for example) to show that $\text{Prob}(X = 0)$ goes to zero as $n \rightarrow \infty$. Then the desired property will be a.a.s. true of $\Gamma \in G(n, p)$ for the given sequence p .

We take a moment to explain a technique that we found expedient in evaluating limits like the one in the above outline. Suppose we want to find $\lim_{n \rightarrow \infty} f(n, p(n))$, where $p(n)$ is a fixed probability sequence and f is some function of two variables. If we partition the sequence $\{(n, p(n))\}$ into two subsequences $\{(n_{a_i}, p(n_{a_i}))\}_i$ and $\{(n_{b_i}, p(n_{b_i}))\}_i$ and

$$\lim_{i \rightarrow \infty} f(n_{a_i}, p(n_{a_i})) = \lim_{i \rightarrow \infty} f(n_{b_i}, p(n_{b_i})) = c \in \mathbb{R} \cup \{-\infty, +\infty\},$$

then of course $\lim_{n \rightarrow \infty} f(n, p(n))$ exists and is the same. Given a sequence $g(n)$, we can choose the subsequences above by choosing $\{n_{a_i}\}$ to be the terms with $p(n) \leq g(n)$ and choosing $\{n_{b_i}\}$ to be the terms with $p(n) > g(n)$, provided that both subsequences are infinite. If one is finite, then we simply ignore it. This allows us to take such a limit in general by dividing it into finitely many cases in which we have added additional hypotheses on $p(n)$.

1.4 Conventions

We always use Γ to denote a finite graph with vertex set V and edge relation \sim . Between functions, \sim denotes asymptotic unity. We use the notation $f \in O(g)$ and $f \lesssim g$ to indicate that eventually f is less than a constant multiple of g , and we also use $O(f)$ to denote an unknown function asymptotically bounded by a constant multiple of f .

2 Domination pairs

2.1 Duality of domination pairs

We will exploit the following connection between adjacent domination and nonadjacent domination. The *link* $\text{lk}(a)$ of a vertex a is $\text{st}(a) \setminus \{a\}$; then $a > b$ if and only if $\text{lk}(b) \subset \text{st}(a)$.

Lemma 2.1 *For $a, b \in V$, we have $a > b$ in Γ if and only if $b > a$ in the complement graph $\bar{\Gamma}$. In particular, $\bar{\Gamma}$ has as many adjacent domination pairs as Γ has nonadjacent ones, and vice versa.*

Proof We add subscripts to our notation for stars and links to make clear which graph we are taking these stars and links in. Of course, $a > b$ in Γ if and only if $\text{lk}_\Gamma(b) \subset \text{st}_\Gamma(a)$. Note that $\text{lk}_{\bar{\Gamma}}(a) = V \setminus \text{st}_\Gamma(a)$, and $\text{st}_{\bar{\Gamma}}(b) = V \setminus \text{lk}_\Gamma(b)$. Then $\text{lk}_{\bar{\Gamma}}(a) \subset \text{st}_{\bar{\Gamma}}(b)$, which proves the lemma. \square

2.2 Existence results

Our existence results follow well-known facts about random graphs. The following statement is taken from Bollobás [1, Theorem III.1, page 57] (the statement in the reference uses some notation from page 5).

Theorem 2.2 *Let $C > 0$ and $\epsilon > 0$ be fixed. If p is a sequence of probabilities such that $p \rightarrow 0$, $p > \epsilon n^{-3/2}$ and*

$$(1) \quad n(n-1)p(1-p)^{n-2} \rightarrow \infty \text{ as } n \rightarrow \infty,$$

then $\Gamma \in G(n, p)$ a.a.s. has at least C vertices of valence 1.

We note a corollary of this:

Corollary 2.3 *If $pn^2 \rightarrow \infty$ and $p < n^{-1}(\log(n) + \log(\log(n)) - \omega(n))$ for some $\omega(n)$ tending to positive infinity, then $\Gamma \in G(n, p)$ a.a.s. has at least C vertices of valence 1 for any fixed C .*

Proof First we suppose that $p(n) > n^{-3/2}$ for all n . Then it is enough to show that the limit in Equation (1) from Theorem 2.2 is satisfied. Since $p \rightarrow 0$, we know $(1 - p)^{n-2} \sim e^{-np}$; so the limit from the theorem is equal to $\lim_{n \rightarrow \infty} n^2 p e^{-np}$. To take this limit, we break the sequence $\{(n, p(n))\}_n$ into two subsequences. The first satisfies $p < (1/4)n^{-1} \log(n)$ and the second satisfies $p \geq (1/4)n^{-1} \log(n)$. For (n, p) in the first subsequence,

$$n^2 p e^{-np} \geq n^{1/2} e^{-\log(n)/4} = n^{1/4} \rightarrow \infty.$$

For (n, p) in the second subsequence,

$$n^2 p e^{-np} \geq \frac{1}{4} n \log(n) e^{-np} = \frac{1}{4} \exp(\log(n) + \log(\log(n)) - np) \rightarrow \infty,$$

since $p < n^{-1}(\log(n) + \log(\log(n)) - \omega(n))$ implies $\log(n) + \log(\log(n)) - np \rightarrow \infty$. Of course, since the limit goes to infinity on both subsequences of $\{(n, p(n))\}_n$, the limit goes to infinity. So if $p > n^{-3/2}$, the theorem implies that Γ a.a.s. has at least a given number of vertices of valence 1.

Next suppose that $p \leq n^{-3/2}$. Let Y_2 denote the random variable that counts the number of vertices of Γ of valence greater than or equal to 2. As explained in the remark preceding Theorem III.1 on page 57 of Bollobás [1],

$$\mathbb{E}(Y_2) \leq \sum_{j=2}^{\infty} \frac{(np)^j}{j!}.$$

Since $np \rightarrow 0$, this sum converges to $e^{np} - np - 1$. Then $\mathbb{E}(Y_2) \rightarrow 0$. By Markov's inequality, we have that $\text{Prob}(Y_2 \geq 1) \rightarrow 0$. So a.a.s., Γ has only vertices of valence 1.

It is a straightforward application of Chebyshev's inequality to show that Γ a.a.s. has at least C edges, but we provide some details for completeness. Let X be the random variable counting the number of edges in Γ . If $X_{a,b}$ is the random variable returning 1 if the distinct vertices a, b in V span an edge and 0 otherwise, then X is the sum of all $X_{a,b}$ over all unordered pairs of vertices. Since $\mathbb{E}(X_{a,b})$ is the probability that a, b span an edge, $\mathbb{E}(X_{a,b}) = p$ and $\mathbb{E}(X) = \binom{n}{2} p$. The second moment of X is

$$\mathbb{E}(X^2) = \binom{n}{2} p + \binom{n}{2} \left(\binom{n}{2} - 1 \right) p^2.$$

This can be computed by expanding $X^2 = (\sum_{a,b} X_{a,b})^2$ and using the fact that edges occur independently. Then

$$\text{Prob}(X < C) \leq \text{Prob}((X - \mathbb{E}(X))^2 > (\mathbb{E}(X) - C)^2) \leq \frac{\mathbb{E}((X - \mathbb{E}(X))^2)}{(\mathbb{E}(X) - C)^2}$$

by Markov’s inequality, and

$$\frac{\mathbb{E}((X - \mathbb{E}(X))^2)}{(\mathbb{E}(X) - C)^2} = \frac{\mathbb{E}(X^2) - \mathbb{E}(X)^2}{\mathbb{E}(X)^2 - 2C\mathbb{E}(X) + C^2} \sim \frac{\mathbb{E}(X^2)}{\mathbb{E}(X)^2} - 1 \rightarrow 0.$$

Since $\text{Prob}(X < C) \rightarrow 0$, Γ a.a.s. has at least C edges. Since Γ a.a.s. has at least C edges and a.a.s. has no vertices of valence greater than one, it follows that Γ a.a.s. has at least $2C$ vertices of valence one (the intersection of two events of probability tending to 1 will have probability tending to 1).

So to prove the corollary in general, we partition the sequence $\{(n, p(n))\}$ into two subsequences, one of which satisfies $p > n^{-3/2}$ and one of which satisfies $p \leq n^{-3/2}$. We have shown that on both subsequences (if they are infinite), the probability that Γ has at least C valence-one vertices goes to one. Of course the same is true for the full sequence. \square

The next statement is from Bollobás [1, Theorem V.4]. An *isolated edge* is one both of whose endpoints have valence 1.

Theorem 2.4 *If $2np - \log(n) - \log(\log(n)) \rightarrow \infty$, then $\Gamma \in G(n, p)$ a.a.s. does not have any isolated edges.*

From these we deduce the following:

Proposition 2.5 *Let $C > 0$ be fixed. If p is in the range*

$$\frac{\log(n) + \log(\log(n)) + \omega_1(n)}{2n} < p < \frac{\log(n) + \log(\log(n)) - \omega_2(n)}{n}$$

for some sequences $\omega_1(n), \omega_2(n)$ approaching positive infinity, then $\Gamma \in G(n, p)$ a.a.s. has at least C vertices of valence 1 that are not on isolated edges.

Proof If Γ has less than C vertices of valence 1 that are not on isolated edges, then either some of the vertices of valence 1 in Γ are on isolated edges, or Γ has less than C vertices of valence 1. Let R denote the event that Γ has less than C vertices of valence 1 that are not on isolated edges, let S denote the event that there are less than C vertices of valence 1, and let T denote the event that there are some isolated edges. Then $R \subset S \cup T$. Any p in the given range certainly satisfies the hypotheses of Theorem 2.4, by the choice of lower bound. Also, p satisfies the hypotheses of Corollary 2.3. So we know $\text{Prob}(S)$ goes to 0 and $\text{Prob}(T)$ goes to 0. Then since

$$0 \leq \text{Prob}(R) \leq \text{Prob}(S) + \text{Prob}(T),$$

we have $\text{Prob}(R) \rightarrow 0$. \square

We need one more classical result, due to Erdős and Rényi [4]. A reference is Bollobás [1, Theorem V.3, Theorem VII.3].

Theorem 2.6 (Erdős–Rényi) *If*

$$p < \frac{\log(n) - \omega(n)}{n}$$

for some $\omega(n) \rightarrow \infty$, then a.a.s. $\Gamma \in G(n, p)$ has at least C isolated vertices for any fixed $C > 0$.

If

$$p > \frac{\log(n) + \omega(n)}{n}$$

for some $\omega(n) \rightarrow \infty$, then a.a.s $\Gamma \in G(n, p)$ is connected.

Proposition 2.7 *Let $C > 0$. If*

$$p < \frac{\log(n) + \log(\log(n)) - \omega(n)}{n}$$

for some sequence $\omega(n)$ with $\omega(n) \rightarrow +\infty$, then a.a.s. there are at least C nonadjacent domination pairs in $\Gamma \in G(n, p)$. If further, $pn^2 \rightarrow \infty$, then there are also at least C adjacent domination pairs. Dually, if

$$1 - p < \frac{\log(n) + \log(\log(n)) - \omega(n)}{n}$$

for some sequence $\omega(n)$ with $\omega(n) \rightarrow +\infty$, then a.a.s. there are at least C adjacent domination pairs in $\Gamma \in G(n, p)$. Finally, if $(1 - p)pn^2 \rightarrow \infty$ as well, then there are also at least C nonadjacent domination pairs.

Proof It is enough to show the first two statements in the proposition, where $p \rightarrow 0$; then the second two statements, where $p \rightarrow 1$, follow dually by [Lemma 2.1](#).

If b is an isolated vertex, then for any other vertex a , the pair (a, b) is a nonadjacent domination pair. If b is a vertex of valence 1 and a is the vertex b is adjacent to, then (a, b) is an adjacent domination pair. If b is a vertex of valence 1 adjacent to a vertex a , and c is some third vertex adjacent to a , then (c, b) is a nonadjacent domination pair. So the number of adjacent domination pairs is at least the number of vertices of valence 1, and the number of nonadjacent domination pairs is at least the number of isolated vertices, plus the number of vertices of valence 1 not on isolated edges.

If p satisfies the more general hypotheses of the proposition, we break it into at most two subsequences, one which satisfies the hypotheses of [Proposition 2.5](#) and the

other which satisfies $np - \log(n) \rightarrow -\infty$. On the first subsequence, the probability of having at least C valence-1 vertices not on isolated edges goes to 1, and on the second subsequence, the probability of having at least C isolated vertices goes to 1 by [Theorem 2.6](#). So the probability of having at least C nonadjacent domination pairs goes to 1 on the entire sequence.

If p satisfies the more restrictive hypotheses, then [Corollary 2.3](#) applies and there are a.a.s. at least C vertices of valence 1. Then a.a.s. we also have at least C adjacent domination pairs. □

2.3 Nonexistence results

We proceed to count nonadjacent domination pairs; our results on adjacent ones follow using [Lemma 2.1](#).

Proposition 2.8 *The expected number of nonadjacent domination pairs in Γ in $G(n, p)$ is*

$$n(n - 1)(1 - p)(p + (1 - p)^2)^{n-2}.$$

Proof Let $X: G(n, p) \rightarrow \mathbb{Z}$ be the random variable with $X(\Gamma)$ equal to the number of pairs (a, b) of distinct vertices in V with a not adjacent to b and $a > b$ in Γ . For each pair $(a, b) \in V^2$ with $a \neq b$, we define a random variable $\hat{X}_{(a,b)}: G(n, p) \rightarrow \mathbb{Z}$ with $\hat{X}_{(a,b)}(\Gamma)$ equal to 1 if a is not adjacent to b and $a > b$, and equal to 0 otherwise. The expectation of $\hat{X}_{(a,b)}$ is the probability that (a, b) is a nonadjacent domination pair.

Suppose $a, b \in V$ with $a \neq b$. For each $c \in V \setminus \{a, b\}$, let $S_c \subset G(n, p)$ denote the event that either $c \sim a$ or both $c \not\sim a$ and $c \not\sim b$. This is a union of two disjoint events whose probabilities are p and $(1 - p)^2$. So $\text{Prob}(S_c) = p + (1 - p)^2$. Since these events involve different edges for different choices of c , the events $\{S_c\}_{c \in V \setminus \{a,b\}}$ are independent. The event $\bigcap_{c \in V \setminus \{a,b\}} S_c$ is exactly the event that every vertex adjacent to b is also adjacent to a , which is by definition the event that $a > b$. So in particular,

$$\text{Prob}(a > b) = \text{Prob}\left(\bigcap_{c \in V \setminus \{a,b\}} S_c\right) = (1 - p + p^2)^{n-2}.$$

The event that a dominates b involves only the edges from a and b to other vertices; in particular, the event that a is nonadjacent to b is independent of it. Then the expectation of $\hat{X}_{(a,b)}$ is the product of the probabilities

$$\mathbb{E}(\hat{X}_{(a,b)}) = (1 - p)(p + (1 - p)^2)^{n-2}.$$

Since X counts the number of nonadjacent domination pairs, and each $X_{(a,b)}$ counts whether (a, b) is such a pair, we have

$$X = \sum_{a \in V} \sum_{b \in V \setminus \{a\}} \hat{X}_{(a,b)}.$$

Then by linearity of expectations, we have

$$\mathbb{E}(X) = n(n-1)(1-p)(p+(1-p)^2)^{n-2}. \quad \square$$

To show that the probability of a domination pair goes to zero in a certain range, we take limits of these expectations and use Markov’s inequality. We will use the following lemma in taking these limits. In fact, counting star-cut-vertices will involve a closely related limit, so to reuse this lemma, we use a parameter k .

Lemma 2.9 *Suppose $k \geq 1$ is an integer and $p = p(n)$ satisfies*

$$2 \frac{\log(n) + \omega(n)}{n} \leq p \leq 1 - (k+1) \frac{\log(n) + \omega(n)}{n}$$

for some sequence $\omega(n)$ that approaches infinity. Let $F(x, y)$ be defined by

$$F(x, y) = x^{k+1} (y + (1-y)^{k+1})^{x-k-1}$$

for suitable $x, y \in \mathbb{R}$. Then $\lim_{n \rightarrow \infty} F(n, p) = 0$.

Proof We take the second partial derivative

$$\begin{aligned} \frac{\partial^2 F}{\partial y^2} &= x^{k+1} (x-k-1) (y + (1-y)^{k+1})^{x-k-3} \\ &\quad \cdot ((x-k-2)(1-(k+1)(1-y)^k)^2 + (k+1)k(1-y)^{k-1}). \end{aligned}$$

For values of y in $[0, 1]$ and values of x in $(k+3, \infty)$, we see that $\partial^2 F / \partial y^2$ is positive, so that F is concave up in its second input (in this range). Let $a(n)$ be the lower bound for p from the statement, and let $1-b(n)$ be the upper bound. Then for large enough n , we have

$$0 \leq F(n, p(n)) \leq \max\{F(n, a(n)), F(n, 1-b(n))\}.$$

Using the well-known bound $(1+s)^t \leq e^{st}$ (for $t > 0$), we see

$$F(n, a(n)) \leq n^{k+1} \exp[(n-k-1)(-1+a(n)+(1-a(n))^{k+1})].$$

We may write this bound as

$$n^{k+1} \exp(-kna(n) + O(na(n)^2))$$

as $n \rightarrow \infty$, using the binomial expansion of $(1 - a(n))^{k+1}$ and the fact that $a(n)$ is $O(na(n)^2)$. This is equivalent to $n^{k+1} \cdot n^{-2k} e^{-2\omega(n)}$ as $n \rightarrow \infty$, so $F(n, a(n)) \rightarrow 0$.

Similarly, we have

$$F(n, 1 - b(n)) \leq n^{k+1} \exp[(n - k - 1)(1 - b(n) + b(n)^{k+1})],$$

which can be written as

$$n^{k+1} \exp(-nb(n) + O(nb(n)^2))$$

as $n \rightarrow \infty$. This is equivalent to $n^{k+1} \cdot n^{-k-1} \exp(-(k + 1)\omega(n))$ as $n \rightarrow \infty$, so $F(n, 1 - b(n)) \rightarrow 0$. Then $F(n, p) \rightarrow 0$ as $n \rightarrow \infty$ for any sequence $p(n)$ in the given range. \square

Proposition 2.10 *Suppose $\omega(n)$ is a sequence of real numbers tending to positive infinity. If $p = p(n)$ is a sequence of probability values satisfying*

$$2 \frac{\log(n) + \omega(n)}{n} \leq p \leq 1 - \frac{\log(n) + \log(\log(n)) + \omega(n)}{n},$$

then the probability that Γ in $G(n, p)$ has a nonadjacent domination pair limits to zero.

Proof Let X be the random variable from the [Proposition 2.8](#). According to that proposition, we have

$$\mathbb{E}(X) \leq n^2(p + (1 - p)^2)^{n-2}.$$

Then by [Lemma 2.9](#), with $k = 1$, we have that $E(X) \rightarrow 0$ as $n \rightarrow \infty$ if we assume that $p \leq 1 - 2n^{-1}(\log(n) + \omega(n))$.

Next we momentarily assume that $p(n) \geq 1 - 3n^{-1} \log(n)$. We change variables to use $q = 1 - p$. By the bound on $\mathbb{E}(X)$ from [Proposition 2.8](#) and the estimate $(1 + s)^t \leq e^{st}$ for $t > 0$, we have

$$\mathbb{E}(X) \leq n^2 q(1 - q + q^2)^{n-2} \leq n^2 q \exp(-nq + O(nq^2)).$$

Since $n^{-1}(\log(n) + \log(\log(n)) + \omega(n)) \leq q \leq 3n^{-1} \log(n)$, we have

$$\mathbb{E}(X) \leq 3n \log(n) \cdot n^{-1} \log(n)^{-1} e^{-\omega(n)} e^{O(n^{-1}(\log(n))^2)}.$$

This certainly limits to zero.

If p satisfies the more general bounds, we subdivide the sequence $(n, p(n))_n$ into two subsequences, the first of which satisfies $p(n) \leq 1 - 2n^{-1}(\log(n) + \omega(n))$, and the second of which satisfies $p(n) \geq 1 - 3n^{-1} \log(n)$. Since $E(X)$ goes to zero on both of these subsequences, we have that $E(X) \rightarrow 0$ as $n \rightarrow \infty$. The proposition follows by Markov’s inequality. \square

To tighten the range of probability functions in which nonadjacent domination pairs occur, we consider the following configuration, which we call a *domination diamond*. This is a quadruple (a, b, c, d) of vertices, all distinct, with $a \sim b \sim c \sim d \sim a$, $a \not\sim c$, $b \not\sim d$ and $a > c$.

Lemma 2.11 *If (a, c) is a nonadjacent domination pair, c is not isolated, and nothing dominates c adjacently, then there is a domination diamond (a, b, c, d) in Γ .*

Proof Of course $\text{lk}(c)$ is nonempty. If the induced subgraph on $\text{lk}(c)$ is a complete graph, then any element of $\text{lk}(c)$ adjacently dominates c . So there are two vertices $b, d \in \text{lk}(c)$ with $b \not\sim d$. Since $b, d \sim c$ and $a > c$, we know $b, d \sim a$. \square

Proposition 2.12 *If $p \rightarrow 0$ and $np \rightarrow \infty$ as $n \rightarrow \infty$, then a.a.s. $\Gamma \in G(n, p)$ has no domination diamonds.*

Proof Let $\widehat{W}_{(a,b,c,d)}$ be the random variable which is 1 if (a, b, c, d) is a domination diamond and 0 otherwise. Then

$$\mathbb{E}(\widehat{W}_{(a,b,c,d)}) = p^4(1 - p)^2(p + (1 - p)^2)^{n-4},$$

since the mandated edges and nonedges among (a, b, c, d) are given, and for any fifth vertex e , we must have e adjacent to c or not adjacent to c or a . Setting W equal to the sum of all $\widehat{W}_{(a,b,c,d)}$ over all choices of (a, b, c, d) , we have that the random variable W counts the number of domination diamonds. By additivity, we have

$$\mathbb{E}(W) = \frac{n!}{(n - 4)!} p^4(1 - p)^2(p + (1 - p)^2)^{n-4}.$$

Since $p \rightarrow 0$, we know

$$(p + (1 - p)^2)^{n-4} \sim \exp(-np).$$

Then

$$\mathbb{E}(W) \sim (np)^4 \exp(-np).$$

Since the function $t \mapsto t^4 e^{-t}$ converges to 0 as $t \rightarrow \infty$, our hypothesis that $np \rightarrow \infty$ forces $\mathbb{E}(W)$ to converge to 0 as $n \rightarrow \infty$. The proposition follows immediately by Markov’s inequality. \square

Theorem 2.13 Suppose $\omega_1(n), \omega_2(n)$ are sequences with $\omega_1(n), \omega_2(n) \rightarrow +\infty$ as $n \rightarrow \infty$. If p satisfies

$$\frac{\log(n) + \log(\log(n)) + \omega_1(n)}{n} < p < 1 - \frac{\log(n) + \log(\log(n)) + \omega_2(n)}{n},$$

then $\Gamma \in G(n, p)$ a.a.s. has no domination pairs.

Proof By Lemma 2.1, it is enough to show that a.a.s. there are no nonadjacent domination pairs. We break p into two subsequences, one where Proposition 2.10 applies, and one where $p \rightarrow 0$. Then it is enough to show that the theorem holds if we assume $p \rightarrow 0$ as $n \rightarrow \infty$.

Let A be the event that there is some adjacent domination pair, B the event that there is some nonadjacent domination pair, C the event that there is some isolated vertex, and D the event that there is some domination diamond. Proposition 2.10 and Lemma 2.1 tell us that $\text{Prob}(A) \rightarrow 0$ in this range, since replacing p with $(1 - p)$ puts the probability function in the range in which there are no nonadjacent domination pairs. Theorem 2.6 implies that $\text{Prob}(C) \rightarrow 0$. Further, Proposition 2.12 tell us that $\text{Prob}(D) \rightarrow 0$.

Lemma 2.11 implies that for every nonadjacent domination pair (a, c) , one of the following holds: (1) c is isolated, (2) there is a domination diamond (a, b, c, d) or (3) something adjacently dominates c . In other words, $B \subset A \cup C \cup D$. Then

$$0 \leq \text{Prob}(B) \leq \text{Prob}(A) + \text{Prob}(C) + \text{Prob}(D)$$

and therefore $\text{Prob}(B) \rightarrow 0$. Of course the theorem then follows from Markov’s inequality. □

3 Star 2–connectedness

3.1 Star separations

A subset $S \subset V$ is a *separation* of Γ if $S \neq \emptyset$, $S \neq V$ and there are no edges in Γ from S to $V \setminus S$. Define a *star separation* of Γ to be a pair (a, S) with $a \in V$ and $S \subset V \setminus \text{st}(a)$, such that S is a separation of $\Gamma \setminus \text{st}(a)$. Call a star separation a star k –separation if $|S| = k$. A star separation (a, S) is *proper* if S is not a separation of Γ . Given a separation S , there is a star separation (a, S) only if there is a vertex $a \in V \setminus S$ such that $\text{st}(a) \neq V \setminus S$. Of course, Γ is star 2–connected if and only if it has no star separations.

If $(a, \{b\})$ is a star separation, then (a, b) is a nonadjacent domination pair. However, the converse does not hold. If $V = \text{st}(a) \sqcup \{b\}$, then (a, b) is a nonadjacent domination pair, but $(a, \{b\})$ is not a star separation. However, this means that in sparse graphs with many vertices, nonadjacent domination pairs are practically the same thing as star 1-separations. This explains the similarities in the functions that describe the expected number of each.

3.2 Counting small star separations

Our first result on star-separations shows that for k not depending on n , star k -separations asymptotically almost certainly do not occur in a wide range of probabilities.

Proposition 3.1 *Let $p = p(n)$ be a sequence of probabilities and let $k \geq 1$ be fixed. Suppose*

$$p \geq \frac{\log(n) + (2/k) \log(\log(n)) + \omega(n)}{n}$$

for some $\omega(n)$ approaching positive infinity. Further suppose that either $k \geq 2$ or else that $n(1-p) \rightarrow 0$ or $n(1-p) \rightarrow \infty$. Then a.a.s. $\Gamma \in G(n, p)$ has no star k -separations.

The proof appears after the next two lemmas. We would like to proceed by fixing k and counting directly the number of star k -separations. However, the random variable that counts star k -separations has a problem: it turns out that there is a range of probabilities where the probability of a star k -separation existing goes to zero even though the expected number of star separations goes to infinity. We get better bounds by counting proper star k -separations instead.

Lemma 3.2 *Let $k > 0$ be an integer and let U_k be the random variable on $G(n, p)$ that counts the number of proper star k -separations. Then the expectation of U_k is*

$$\mathbb{E}(U_k) = n \binom{n-1}{k} (1-p)^k \cdot [(p + (1-p)^{k+1})^{n-k-1} + (1-p^{n-k-1})(1-(1-p)^{k(n-k-1)}) - 1].$$

Proof Let U_k count the number of proper star k -separations (a, S) for various $a \in V$ and $S \subset V \setminus \{a\}$ (with $|S| = k$). Let $\hat{U}_{(a,S)}$ be the random variable with value 1 if (a, S) is a proper star k -separation and value 0 otherwise.

The expectation $\mathbb{E}(\hat{U}_{(a,S)})$ is the probability that (a, S) is a proper star k -separation. If (a, S) is a proper star separation, then necessarily a is not adjacent to any element of S . This event has probability $(1-p)^k$. This event is independent of the other

aspects of the definition that we are about to describe, since we will not mention these edges again.

The next feature of the definition is that we must have every element of $V \setminus (S \cup \{a\})$ either in $\text{lk}(a)$ or not adjacent to any member of S . Suppose $b \in V \setminus (S \cup \{a\})$. The event that b is adjacent to a has probability p ; the event that b is not adjacent to any element of $S \cup \{a\}$ is disjoint from this event and has probability $(1 - p)^{k+1}$. So the probability that b is either adjacent to a or not adjacent to any member of S has probability $p + (1 - p)^{k+1}$. Since these events involve different edges for different choices of b , they are independent, and the event that every element of $V \setminus (S \cup \{a\})$ is in $\text{lk}(a)$ or not adjacent to anything in S has probability

$$(p + (1 - p)^{k+1})^{n-k-1}.$$

Next, we must exclude two events that are included in the previous event. To fit the definition of a star separation, we must have that $V \setminus S$ is not all of $\text{st}(a)$. This means that we are excluding the event that a is adjacent to every element of $V \setminus S$, which has probability p^{n-k-1} . To ensure that (S, a) is a proper star separation, we must exclude the event that S is a separation. Since we have already stipulated that there are no edges from a to S , this is then the event that there are no edges from $V \setminus (S \cup \{a\})$ to S . This has probability $(1 - p)^{k(n-k-1)}$. The two events we have just described are independent, so by DeMorgan's law, the probability of either event happening is

$$1 - (1 - p^{n-k-1})(1 - (1 - p)^{n(n-k-1)}).$$

Putting this all together, we see that

$$\mathbb{E}(\widehat{U}_{(a,S)}) = (1-p)^k [(p+(1-p)^{k+1})^{n-k-1} + (1-p^{n-k-1})(1-(1-p)^{k(n-k-1)}) - 1].$$

We note that there are n possible choices for a in V , and given a choice of a , there are $\binom{n-1}{k}$ choices for S . Since U_k is the sum of $\widehat{U}_{(a,S)}$ over all choices of (a, S) , the result follows from the linearity of expectations. □

Next we process this expression into a pair of more manageable bounds.

Lemma 3.3 *We have the following bounds on $\mathbb{E}(U_k)$:*

$$\begin{aligned} \mathbb{E}(U_k) &\leq n^2 \binom{n-1}{k} (1-p)^{2k+1} (p + (1-p)^{k+1})^{n-k-2}, \\ \mathbb{E}(U_k) &\leq kn^2 \binom{n-1}{k} (1-p)^{k+1} p^2 (p + (1-p)^{k+1})^{n-k-2}. \end{aligned}$$

Proof These bounds come from the expression in [Lemma 3.2](#) in essentially the same way. We will use the following claim both times.

Claim *Suppose a and b are real numbers with $0 \leq a \leq b \leq 1$ and n is a natural number. Then*

$$a^n - b^n \leq na^{n-1}(b - a).$$

The claim follows easily from a calculus argument: if

$$f(a) = na^{n-1}(b - a) + b^n - a^n,$$

then $f(b) = 0$, but $f'(a)$ is negative for a in $(0, b)$.

Now using the claim, we deduce the lemma from [Lemma 3.2](#). Since $1 - p^{n-k-1} \leq 1$, we deduce first that

$$\mathbb{E}(U_k) \leq n \binom{n-1}{k} (1-p)^k ((p + (1-p)^{k+1})^{n-k-1} - (1-p)^{k(n-k-1)}).$$

Since $0 \leq (1-p)^k \leq p + (1-p)^{k+1} \leq 1$, the claim implies

$$\begin{aligned} (p + (1-p)^{k+1})^{n-k-1} - (1-p)^{k(n-k-1)} &\leq (n-k-1)(p + (1-p)^{k+1})^{n-k-2}(p + (1-p)^{k+1} - (1-p)^k) \\ &= (n-k-1)(p + (1-p)^{k+1})^{n-k-2} p(1 - (1-p)^k). \end{aligned}$$

Since $(1-p)^k \geq 1 - kp$ for $p \in [0, 1]$, it follows that

$$\mathbb{E}(U_k) \leq kn(n-k-1) \binom{n-1}{k} p^2 (1-p)^{k+1} (p + (1-p)^{k+1})^{n-k-2}.$$

Second, since $1 - (1-p)^{k(n-k-1)} < 1$, we have

$$\mathbb{E}(U_k) \leq n \binom{n-1}{k} (1-p)^k ((p + (1-p)^{k+1})^{n-k-1} - p^{n-k-1}).$$

Then since $0 \leq p \leq p + (1-p)^{k+1} \leq 1$, the claim implies

$$\mathbb{E}(U_k) \leq n(n-k-1) \binom{n-1}{k} (1-p)^{2k+1} (p + (1-p)^{k+1})^{n-k-1}.$$

The lemma follows. □

Proof of Proposition 3.1 As usual, we split $\{(n, p(n))\}_n$ into subsequences satisfying stronger hypotheses that overlap. Therefore it is enough to show the proposition for p in several subcases with stronger hypotheses. First we suppose that p satisfies

$$2 \frac{\log(n) + \omega(n)}{n} \leq p \leq 1 - (k + 1) \frac{\log(n) + \omega(n)}{n}.$$

We get the following bound from Lemma 3.2:

$$\mathbb{E}(U_k) \leq n^{k+1} (p + (1 - p)^{k+1})^{n-k-1}$$

(This is because $(1 - p^{n-k-1})(1 - (1 - p)^{k(n-k-1)}) - 1 \geq 0$.) Then by Lemma 2.9, we have that $\mathbb{E}(U_k)$ converges to 0 for any p in this range.

Next we suppose that $q \rightarrow 0$ and $nq^2 \rightarrow 0$, where $q = 1 - p$. By the first bound from Lemma 3.3, we see that

$$\mathbb{E}(U_k) \leq n^{k+2} q^{2k+1} (1 - q + q^{k+1})^{n-k-2}.$$

Since $(1 - x)^m \sim e^{-mx}$ as $x \rightarrow 0$, this bound is asymptotically equivalent to

$$n^{k+2} q^{2k+1} \exp(-nq + O(nq^2) + O(q)),$$

which is equivalent to

$$q^{k-1} (nq)^{k+2} e^{-nq}.$$

From calculus, we know that $x \mapsto x^{k+2} e^{-x}$ is bounded and tends to zero as $x \rightarrow \infty$ or $x \rightarrow 0$. So if $k > 1$ or $nq \rightarrow 0$ or $nq \rightarrow \infty$, we know that $\mathbb{E}(U_k) \rightarrow 0$.

Finally, we suppose that $p \rightarrow 0$, $np^2 \rightarrow 0$ and

$$p > \frac{\log(n) + (2/k) \log(\log(n)) + \omega(n)}{n}$$

for some $\omega(n) \rightarrow \infty$. Using the second bound from Lemma 3.3, we have

$$\mathbb{E}(U_k) \leq n^{k+2} p^2 (p + (1 - p)^{k+1})^{n-k-2}.$$

By binomial expansion, this bound can be written as

$$n^{k+2} p^2 (1 - kp + O(p^2))^{n-k-2}.$$

This is asymptotically equivalent to

$$n^{k+2} p^2 \exp(-knp + O(np^2) + O(p)) \sim n^{k+2} p^2 e^{-knp},$$

since $np^2, p \rightarrow 0$. Then by our lower bound on p , we have

$$\begin{aligned} \mathbb{E}(U_k) &\lesssim n^{k+2} \cdot n^{-2}(\log(n) + \log(\log(n)) + \omega(n))^2 \cdot n^{-k} \log(n)^{-2} e^{-k\omega(n)} \\ &= \left(\frac{\log(n) + \log(\log(n)) + \omega(n)}{\log(n)} \right)^2 e^{-k\omega(n)} \lesssim \omega(n)^2 e^{-k\omega(n)}. \end{aligned}$$

So $\mathbb{E}(U_k) \rightarrow 0$.

We have shown under these hypotheses there are a.a.s. no proper star k -separations. However, [Theorem 2.6](#) implies that a.a.s. there are no separations for p in this range. So a.a.s. there are no star k -separations. □

3.3 Summed counts of star separations

So far, we have shown that for each k , $\mathbb{E}(U_k) \rightarrow 0$ in a certain range of probability values. We would like to show that $\mathbb{E}(\sum_k U_k) \rightarrow 0$ for p in a specific range. Effectively, this requires commuting a limit and sum, which the Lebesgue dominated convergence theorem would allow. To meet the hypotheses of this theorem, we must compute bounds on $\mathbb{E}(U_k)$. We do this using calculus techniques, using two different bounds that are useful when p is close to 1 and when p is close to 0, respectively. Unfortunately, showing these bounds carefully takes a fair amount of work. We proceed to prove [Theorem 1.5](#), and then we prove the bounds on $\mathbb{E}(U_k)$.

Proof of Theorem 1.5 We break the sequence $\{(n, p(n))\}_n$ into two subsequences, where one satisfies $p \leq 2/5$, and the other satisfies $p \geq 2/5$. We prove the theorem for each subsequence; then it follows that the theorem is true for the original sequence.

Suppose the first hypothesis of the theorem, that $p > n^{-1}(\log(n) + \log(\log(n)) + \omega(n))$. We note that every star 1-separation consists of a nonadjacent domination pair. Then [Theorem 2.13](#) shows that a.a.s., there are no star 1-separations if $p \leq 2/5$. If $p \geq 2/5$, then [Proposition 3.1](#) includes the fact that there are a.a.s. no star 1-separations, provided that the second hypothesis also holds. Then quoting [Proposition 3.1](#), we have

$$\lim_{n \rightarrow \infty} \mathbb{E}(U_k) = 0,$$

for any fixed k , if $k \geq 2$ (and p satisfies the first hypothesis) or if $k = 1$ and p satisfies both hypotheses of the theorem.

[Proposition 3.5](#) below states that if $p \geq 2/5$, then there is a nonnegative sequence $\{a_k\}_k$ with $\mathbb{E}(U_k) \leq a_k$ for all n and for all k with $2k \leq n$, and such that $\sum a_k < \infty$. [Proposition 3.6](#) is the same statement in the case that $p \leq 2/5$. Then in either case,

the Lebesgue dominated convergence theorem (see, eg Rudin [6, page 26]) applies. Therefore if we assume both hypotheses, we have

$$(2) \quad \lim_{n \rightarrow \infty} \sum_{k=1}^{\lfloor n/2 \rfloor} \mathbb{E}(U_k) = \sum_{k=1}^{\infty} \lim_{n \rightarrow \infty} \mathbb{E}(U_k) = 0.$$

Of course, by linearity of expectations, we have

$$\sum_{k=1}^{\lfloor n/2 \rfloor} \mathbb{E}(U_k) = \mathbb{E} \left(\sum_{k=1}^{\lfloor n/2 \rfloor} U_k \right).$$

The random variable on the right will be zero only if $\Gamma \in G(n, p)$ has no proper star-separations: if $a \in \Gamma$ is a star-cut-vertex, then some component of $\Gamma \setminus \text{st}(a)$ has less than $n/2$ vertices. Then by Equation (2) there are a.a.s. no proper star separations. Since Theorem 2.6 implies there are a.a.s. no separations, we know there are a.a.s. no star-cut-vertices.

Similarly, if we assume only the first hypothesis, Equation (2) will be true if we sum from $k = 2$ to $k = \lfloor n/2 \rfloor$ on both sides, instead of starting at $k = 1$. Then we deduce that

$$\lim_{n \rightarrow \infty} \mathbb{E} \left(\sum_{k=2}^{\lfloor n/2 \rfloor} U_k \right) = 0.$$

This random variable will be zero only if there are no proper star k -separations for $2 \leq k < n/2$. If the star of a star-cut-vertex has more than one complementary component with at least two vertices, then it will have a complementary component with at least two and fewer than $n/2$ vertices. The second statement in the theorem follows. □

To prove our bounds on $\mathbb{E}(U_k)$, we first prove a bound on the binomial coefficients for which we were unable to find a reference.

Lemma 3.4 *For any integers n, k with $n - 1 \geq k \geq 0$, we have*

$$\binom{n-1}{k} \leq \frac{n^n}{(n-k)^{n-k} k^k}.$$

Proof Using the product expression for the binomial coefficient, we have

$$\log \left(\binom{n-1}{k} \right) = \sum_{i=n-k}^{n-1} \log(i) - \sum_{i=2}^k \log(i).$$

We recognize $\sum_{i=n-k}^{n-1} \log(i)$ as a lower Riemann sum approximation to the integral $\int_{n-k}^n \log(t) dt$, so that

$$\sum_{i=n-k}^{n-1} \log(i) \leq n \log(n) - (n-k) \log(n-k) - k.$$

Similarly, $\sum_{i=2}^k \log(i)$ is an upper Riemann sum approximation to $\int_1^k \log(t) dt$, so

$$\sum_{i=1}^k \log(i) \geq k \log(k) - k + 1.$$

Then putting these together,

$$\log\left(\binom{n-1}{k}\right) \leq n \log(n) - (n-k) \log(n-k) - k \log(k) - 1.$$

Exponentiating,

$$\binom{n-1}{k} \leq \frac{1}{e} \cdot \frac{n^n}{(n-k)^{n-k} k^k},$$

and the lemma follows. □

Now we bound $\mathbb{E}(U_k)$. The choice of $2/5$ below is somewhat arbitrary.

Proposition 3.5 *There is a sequence of positive numbers a_k such that $\sum_{k=1}^\infty a_k < \infty$ and $\mathbb{E}(U_k) \leq a_k$ for any $n \geq 2k$ and any $p \geq 2/5$.*

Proof We write the first bound on $\mathbb{E}(U_k)$ from [Lemma 3.3](#) in terms of $q = 1 - p$:

$$\mathbb{E}(U_k) \leq n \binom{n-1}{k} (n-k-1) q^{2k+1} (1-q+q^{k+1})^{n-k-2}.$$

In terms of q , our hypothesis is that $q \leq 3/5$. We set

$$F(k, n, q) = \frac{n^n}{(n-k)^{n-k} k^k} n^2 q^{2k+1} (1-q+q^{k+1})^{n-k-2};$$

then by [Lemma 3.4](#),

$$\mathbb{E}(U_k) \leq F(k, n, q).$$

We find the bounding sequence by using vector calculus to find critical points for $F(k, n, q)$ for fixed k and for (n, q) in the region $[2k, \infty) \times [0, 3/5]$. In the following, we treat n and q as independent variables.

Claim *For large enough k , the partial derivative $\partial \log \circ F / \partial n$ is never zero on the vertical ray $(n, q) \in [2k, \infty) \times \{3/5\}$.*

To show this, we consider

$$\frac{\partial \log \circ F}{\partial n}(k, n, q) = \frac{2}{n} + \log\left(\frac{n}{n-k}\right) + \log(1 - q + q^{k+1}).$$

This is zero if and only if

$$e^{-2/n}(n - k) + n(-1 + q - q^{k+1}) = 0.$$

Define $f_k(n, q) = e^{-2/n}(n - k) + n(-1 + q - q^{k+1})$. Next we define $g_k(n, q) = -(k + 2) + n(q - q^{k+1})$. We use the following bound from calculus:

$$|e^{-2/n} - (1 - 2/n)| \leq 2/n^2.$$

This implies that

$$f_k(n, q) - g_k(n, q) \leq 2/n + k(2/n^2 - 2/n) \leq 2/k,$$

since $n \geq 2k$.

The function $g_k(n, q)$ is chosen so that we can solve $g_k(n, q) = 0$ for n easily:

$$g_k(n, q) = 0 \quad \text{if and only if} \quad n = \frac{k + 2}{q - q^{k+1}}.$$

To show the claim, we show that $f_k(n, q)$ is not zero on the vertical ray. Note that $g_k(n, 3/5) \geq -(k + 2) + 2k((3/5) - (3/5)^{k+1})$ on this ray. Since this bound is asymptotically equivalent to $(1/5)k$ we see that for large k , $g_k(n, 3/5) > 2/k$ on the ray. This implies that $f_k(n, 3/5)$ is not zero when $n > 2k$ and k is large enough, proving the claim.

Claim *The function $\log \circ F$ never has zero gradient on $[2k, \infty) \times [0, 3/5]$ for large enough k .*

Note that

$$(3) \quad \frac{\partial \log \circ F}{\partial q}(k, n, q) = \frac{2k + 1}{q} + (n - k - 2) \frac{-1 + (k + 1)q^k}{1 - q + q^{k+1}}.$$

We suppose that this partial derivative is 0 and $f_k(n, q)$ is zero as well. If $f_k(n, q)$ is zero, then $g_k(n, q) = \epsilon$ for some $\epsilon = \epsilon(n, q, k)$ with $|\epsilon| > 2/k$. Then

$$q - q^{k+1} = \frac{k + 2 - \epsilon}{n}.$$

Assuming that $(\partial \log \circ F)/\partial q = 0$, then substituting this for one instance of $q - q^{k+1}$ in the expression in Equation (3), we get the equation

$$\frac{2k + 1}{q} + \frac{n - k - 2}{n - k - 2 + \epsilon} \cdot n(-1 + (k + 1)q^k) = 0.$$

Next we use the substitution $n = (k + 2 - \epsilon)/(q - q^{k+1})$ on one of the instances of n to get

$$\frac{2k + 1}{q} + \frac{n - k - 2}{n - k - 2 + \epsilon} \cdot \frac{k + 2 - \epsilon}{q - q^{k+1}} \cdot (-1 + (k + 1)q^k) = 0.$$

We write $c = (n - k - 2)/(n - k - 2 + \epsilon)$. Note that c is close to 1. From here, it is straightforward to solve for q^k :

$$q^k = \frac{(2 - c)k + 1 - c(2 - \epsilon)}{-ck^2 + (2 - c(3 - \epsilon))k + 1 - c(2 - \epsilon)}.$$

The right side of this equation is negative, so it has no real solution for q if k is even. Taking the limit as $k \rightarrow \infty$ on odd k , we see that $q \rightarrow -1$. This implies that for large enough k , we never have both $f_k(n, q) = 0$ and $(\partial \log \circ F)/\partial q = 0$; this proves the claim.

Claim *For large enough k , the maximum of $F(k, n, q)$ for $(n, q) \in [2k, \infty) \times [0, 3/5]$ is realized at $(n, q) = (2k, 3/5)$.*

We have shown that the gradient of $\log \circ F$ never vanishes on $[2k, \infty) \times [0, 3/5]$, and that the partial with respect to n never vanishes on the left boundary of the region. Note that F is zero on the right boundary. So it is enough to show that F is increasing in q for $q < 3/5$ when $n = 2k$. It is straightforward to see from Equation (3) that

$$\frac{\partial \log \circ F}{\partial q} \geq \frac{2k}{q} - \frac{k - 2}{1 - q}.$$

However, $q \mapsto 2k/q - (k - 2)/(1 - q)$ is plainly decreasing in q ; evaluating it at $3/5$, we see that $(\partial \log \circ F)/\partial q \geq (5/6)k + 5 > 0$. This proves this last claim.

Then to prove the lemma, we note

$$\lim_{k \rightarrow \infty} \frac{F(k + 1, 2k + 2, 3/5)}{F(k, 2k, 3/5)} = \frac{72}{125} < 1.$$

So the maximum value of $\mathbb{E}(U_k)$ for q in this range is eventually bounded above by an exponentially decaying function of k . □

Proposition 3.6 *There is a sequence of positive numbers a_k such that $\sum_{k=1}^{\infty} a_k < \infty$ and $\mathbb{E}(U_k) \leq a_k$ for any $n \geq 2k$ and any p satisfying*

$$\frac{\log(n) + \log(\log(n))}{n} \leq p \leq 2/5.$$

Proof Define $G(k, n, p)$ by

$$G(k, n, p) = \frac{n^n}{(n-k)^{n-k} k^{k-1}} n^2 p^2 (1-p)^{k+1} (p + (1-p)^{k+1})^{n-k-2}.$$

Then by [Lemma 3.3](#) (the second bound) and [Lemma 3.4](#), we know that

$$G(k, n, p) \geq \mathbb{E}(U_k).$$

To bound $\mathbb{E}(U_k)$ by a function of k , we will show that for sufficiently large fixed values of k , the maximum of G on the region

$$R = \left\{ (n, p) \mid n \geq 2k, \frac{\log(n) + \log(\log(n))}{n} \leq p \leq 2/5 \right\}$$

is at one of the corners of R .

Claim 1 *For large enough k , the gradient of $\log \circ G$ is never zero on R .*

Since $p \mapsto 1 - p - (1-p)^{k+1}$ is concave down, it is easy to verify that for large enough k and $p < 2/5$, we have

$$1 - p - (1-p)^{k+1} \geq \min\left(\frac{(k+2)p}{2}, \frac{k+2}{2k}\right).$$

We use this inequality in a bound on

$$\frac{\partial \log \circ G}{\partial n}(k, n, p) = \frac{2}{n} + \log\left(1 + \frac{k}{n-k}\right) + \log(p + (1-p)^{k+1}).$$

It is immediate that

$$\frac{\partial \log \circ G}{\partial n}(k, n, p) \leq \frac{2}{n} + \frac{k}{n-k} - 1 + p + (1-p)^{k+1},$$

from which we deduce

$$\frac{\partial \log \circ G}{\partial n}(k, n, p) \leq \frac{2}{n} + \frac{k}{n-k} - \min\left(\frac{(k+2)p}{2}, \frac{k+2}{2k}\right).$$

If $(k+2)/(2k)$ is the smaller quantity, then it is straightforward to check that this bound is negative (using $n \geq 2k$). If $((k+2)p)/2$ is the smaller quantity, then using $n \geq 2k$ and $p \geq n^{-1}(\log(n) + \log(\log(n)))$, it is also routine to check that the bound is

negative for large k . Then $(\partial \log \circ G)/\partial n$ is always negative on R , so that the gradient is never zero on R .

Next we deal with the sloping boundary of R .

Claim 2 For all k large enough and $n \geq 2k$, the function $G(k, n, p(n))$ is decreasing in n , where

$$p = p(n) = \frac{\log(n) + \log(\log(n))}{n}.$$

First we compute the logarithmic partial derivative of $G(k, n, p(n))$ with respect to n :

$$\begin{aligned} \frac{\partial}{\partial n} \log(G(k, n, p(n))) &= \log\left(\frac{n}{n-k}\right) + \frac{2}{n} + \frac{2p'}{p} - \frac{kp'}{1-p} + \log(p + (1-p)^{k+1}) \\ &\quad + p'(n-k-2) \frac{1-(k+1)(1-p)^k}{p + (1-p)^{k+1}} \end{aligned}$$

In computing upper bounds on this expression, we will freely assume that k is large, and we will always assume that $n \geq 2k$.

Since $1 - (k + 1)^{-1/k}$ goes to 0 more slowly than $p(2k)$ as $k \rightarrow \infty$, we may assume that $1 - (k + 1)(1 - p)^k$ is negative. We may assume that p' is negative. Further, note that $p + (1 - p)^{k+1} \geq (1 - p)^k$. Then

$$p'(n-k-2) \frac{1-(k+1)(1-p)^k}{p + (1-p)^{k+1}} \leq (n-k-2)p'((1-p)^{-k} - 1 - k).$$

By logarithm rules and the inequality $\log(1 + x) \leq x$, we may deduce

$$\log\left(\frac{n}{n-k}\right) \leq \frac{k}{n-k} \quad \text{and} \quad \log(p + (1-p)^{k+1}) \leq p((1-p)^{-k-1} - 1 - k).$$

Then we have

$$\begin{aligned} (4) \quad \frac{\partial}{\partial n} \log(G(k, n, p(n))) &\leq \frac{k}{n-k} + \frac{2}{n} + \frac{2p'}{p} - \frac{kp'}{1-p} + p((1-p)^{-k-1} - 1 - k) \\ &\quad + (n-k-2)p'((1-p)^{-k} - 1 - k). \end{aligned}$$

To process this expression, we start combining terms. First of all, it is straightforward to show

$$\frac{2}{n} + \frac{2p'}{p} = 2 \frac{\log(n) + 1}{\log(n)(\log(n) + \log(\log(n)))} \leq \frac{2}{\log(n)}.$$

Next, we note

$$\frac{k}{n-k} - k(p + np') = \frac{k^2}{n(n-k)} - \frac{k}{n \log(n)}.$$

Since

$$p' \leq -\frac{\log(n)}{n^2},$$

we have

$$\frac{k}{n-k} - k(p + (n-k-2)p') \leq \frac{k^2}{n(n-k)} - \frac{k}{n \log(n)} - \frac{k(k+2) \log(n)}{n^2}.$$

Next we note that

$$\frac{-kp'}{1-p} \leq \frac{2k \log(n)}{n^2}.$$

Since the $-k(k+2) \log(n)/n^2$ dominates the positive terms, we may deduce that for any positive constant strictly less than 1, say $1/2$, we have

$$(5) \quad \frac{k}{n-k} - k(p + (n-k-2)p') + \frac{2}{n} + \frac{2p'}{p} - \frac{kp'}{1-p} \leq -\frac{k}{n \log(n)} - \left(\frac{1}{2}\right) \frac{k^2 \log(n)}{n^2}.$$

The remaining terms in our bound from Equation (4) may be written as

$$(6) \quad (p + (n-k-2)p')((1-p)^{-k-1} - 1) - pp'(n-k-2)(1-p)^{-k-1}.$$

Since we assume that $p < 1/2$, we know $-2 \log(2)p < \log(1-p)$. Since $(1-p)^{-k-1}$ is decreasing in n for fixed k , we may get an upper bound by plugging in for $n = 2k$. Then

$$(1-p)^{-k-1} \leq (2k \log(2k))^{\log(2)(k+1)/k} \leq 2k^{3/4}.$$

This gives us an immediate upper bound on the second term from Equation (6) as follows, using the obvious bound $p \leq 2 \log(n)/n$:

$$-pp'(n-k-2)(1-p)^{-k-1} \leq \frac{8k^{3/4}(\log(n))^2}{n^2}.$$

Next we bound the first term in Equation (6). Since $(1-p)^{k+1} \geq 1 - (k+1)p$, we deduce that

$$(1-p)^{-k-1} - 1 \leq (k+1)p(1-p)^{-k-1} \leq \frac{4(k+1)k^{3/4} \log(n)}{n}.$$

Then

$$(p + (n-k-2)p')((1-p)^{-k-1} - 1) \leq \frac{8(k+1)k^{3/4} \log(n)}{n^2} + \frac{8(k+1)(k+2)k^{3/4} \log(n)^2}{n^3}.$$

Using our assumption that $n \geq 2k$, we may then deduce that

$$(p + (n - k - 2)p')((1 - p)^{-k-1} - 1) \leq \frac{8(k + 1)k^{3/4} \log(n)}{n^2} + \frac{8(k + 1) \log(k)k^{3/4} \log(n)}{n^2}.$$

Then the $-k^2 \log(n)/n^2$ term from Equation (5) eventually (for large k and all $n \geq 2k$) dominates everything from Equation (6), and the claim follows.

Claim 3 *For large fixed k , the minimum value of $G(k, n, p)$ in R is realized at one of the corners of R :*

$$(n, p) = (2k, 2/5) \quad \text{or} \quad (n, p) = (2k, (2k)^{-1}(\log(2k) + \log(\log(2k))))).$$

We have shown that the gradient of $\log \circ G$ never vanishes on R and G is decreasing on the sloping boundary of R . In showing that the gradient never vanishes, we showed that $(\partial \log \circ G)/\partial n$ is negative on the boundary segment $p = 2/5$. This means that the maximum cannot occur on this boundary segment (away from the corner). Now we consider the boundary segment where $n = 2k$. We consider

$$\frac{\partial \log \circ G}{\partial p} = \frac{2}{p} - \frac{k}{1 - p} + (n - k - 2) \frac{1 - (k + 1)(1 - p)^k}{p + (1 - p)^{k+1}}.$$

When $n = 2k$, this is the function

$$p \mapsto \frac{2}{p} - \frac{k}{1 - p} + (k - 2) \frac{1 - (k + 1)(1 - p)^k}{p + (1 - p)^{k+1}},$$

where $(2k)^{-1}(\log(2k) + \log(\log(2k))) \leq p \leq 2/5$. We define a function

$$f: ((2k)^{-1}(\log(2k) + \log(\log(2k))), 2/5) \rightarrow \mathbb{R}$$

to be this function with denominators cleared:

$$f(p) = (2(1 - p) - kp)(p + (1 - p)^{k+1}) + (k - 2)p(1 - p)(1 - (k + 1)(1 - p)^k).$$

Computations show

$$f''(p) = -4k + (1 - p)^{k-2}g(k, p),$$

where g is some polynomial in k and p and

$$f'''(p) = (1 - p)^{k-3}(2k - 5k^3 - 3k^4 + (-2k + 7k^3 + 6k^4 + k^5)p - (2k^3 + 3k^4 + k^5)p^2).$$

A computation shows that $f'''(p)$ is positive if

$$\frac{-2 + 2k + 3k^2}{k^2(2 + k)} < p < 1,$$

so for large k , $f'''(p)$ is positive on the entire domain of f . Then $f''(p) \leq f''(2/5)$, for p in the domain of f . It is easy to see that $f''(2/5) \sim -4k$ as $k \rightarrow \infty$, so that $f''(2/5)$ is negative for large k and p in the domain of f . Then f is concave down. Since $f(2/5) \sim \frac{2}{25}k$ as $k \rightarrow \infty$, this means that f changes sign at most once. Then G is decreasing then increasing as p increases along the boundary segment of R with $n = 2k$. In particular, this proves the claim.

Finally, we can prove the proposition. We note that

$$\lim_{k \rightarrow \infty} \frac{G(k + 1, 2k + 2, 2/5)}{G(k, 2k, 2/5)} = 24/25 < 1,$$

$$\lim_{k \rightarrow \infty} \frac{G(k + 1, 2k + 2, p(2k + 2))}{G(k, 2k, p(2k))} = 0,$$

where $p(n) = n^{-1}(\log(n) + \log(\log(n)))$. Then the sequence of values of G at each corner of R is eventually dominated by an exponentially decreasing function. \square

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