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DISTRIBUTION ESTIMATES OF BARRIER-CROSSING PROBABILITIES OF THE YEH-WIENER PROCESS

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Let $Q = [0, S] \times [0, T]$ be a rectangle and $\{X(s, t); s, t \geq 0\}$ be the two-parameter Yeh-Wiener process. This paper finds probabilities of $X(s, t)$ crossing barriers of the type $ast + bs + ct + d$ on the boundary ∂Q . These probabilities give lower bounds for the yet unknown probabilities of $X(s, t)$ crossing $ast + bs + ct + d$ on Q . The paper also discusses sharper bounds for the latter probabilities.

1. Introduction. Let $\{X(s, t); s, t \geq 0\}$ be the standard Yeh-Wiener process of two parameters such that it is a separable real Gaussian stochastic process satisfying:

$$(1.1) \quad X(s, t) = 0 \text{ a.s. if } s \text{ or } t \text{ is } 0,$$

$$(1.2) \quad \text{the expected value } E\{X(s, t)\} = 0 \text{ at every } s, t \geq 0,$$

$$(1.3) \quad E\{X(s, t)X(s', t')\} = \min(s, s') \cdot \min(t, t').$$

Further properties of the process are found in Yeh's [8] and [9].

For the square $D = [0, 1] \times [0, 1]$ and its boundary ∂D , Paranjape and Park [6] showed that the probability

$$(1.4) \quad P\left\{\sup_{\partial D} X(s, t) \geq \lambda\right\} = 3N(-\lambda) - e^{4\lambda^2}N(-3\lambda), \quad \lambda \geq 0,$$

where $N(\cdot)$ stands for the standard normal distribution function. This probability is a lower bound of the yet unknown probability, $P\{\sup_D X(s, t) \geq \lambda\}$. It is known (see [4] or [7]) that

$$(1.5) \quad P\left\{\sup_D X(s, t) \geq \lambda\right\} \leq 4P\{X(1, 1) \geq \lambda\} = 4N(-\lambda).$$

Recently Chan [1] showed that, for every $\varepsilon > 0$,

$$(1.6) \quad P\left\{\sup_D X(s, t) \geq \lambda\right\} \leq N(\varepsilon)^{-1}P\left\{\sup_D X(s, t) \geq \lambda - \varepsilon\right\}.$$

By the same technique as he used in his paper, the upper bound can easily be improved to $N(\varepsilon)^{-1}P\{\sup X(1, t) \geq \lambda - \varepsilon; 0 \leq t \leq 1\} = 2N(-\lambda + \varepsilon)/N(\varepsilon)$. However it turns out to be that even this improved upper bound is not as good as $4N(-\lambda)$ for any $\varepsilon > 0$. In fact

$$4N(-\lambda) < N(\varepsilon)^{-1}P\left\{\sup_{0 \leq t \leq 1} X(1, t) \geq \lambda - \varepsilon\right\}, \quad \varepsilon > 0,$$

and

$$\lim_{\varepsilon \rightarrow 0^+} N(\varepsilon)^{-1} P \left\{ \sup_{0 \leq t \leq 1} X(1, t) \geq \lambda - \varepsilon \right\} = 4N(-\lambda).$$

More recently Goodman [3] showed that for $\lambda \geq 0$,

$$(1.7) \quad 2 \left\{ N(-\lambda) + \lambda \int_{\lambda}^{\infty} N(-s) ds \right\} \leq P \left\{ \sup_D X(s, t) \geq \lambda \right\}.$$

Obviously the left-hand side of (1.7) is a much better lower bound of $P\{\sup_D X(s, t) \geq \lambda\}$ than (1.4). He subsequently proves that

$$(1.8) \quad \lim_{\lambda \rightarrow \infty} \frac{2 \{ N(-\lambda) + \lambda \int_{\lambda}^{\infty} N(-s) ds \}}{4N(-\lambda)} = 1,$$

thus showing that both $2 \left\{ N(-\lambda) + \lambda \int_{\lambda}^{\infty} N(-s) ds \right\}$ and $4N(-\lambda)$ are very good approximations of $P\{\sup_D X(s, t) \geq \lambda\}$ for all sufficiently large λ .

The main purpose of this paper is to generalize the above results for more general barriers, namely, to find a formula for

$$P \left\{ \sup_{\partial D} X(s, t) - (ast + bs + ct + d) \geq 0 \right\}, \quad a, b, c, d \geq 0,$$

and then find a lower bound for $P\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\}$ for which (1.7) is a special case. It is apparent that for all $a, b, c, d \geq 0$

$$(1.9) \quad P \left\{ \sup_D X(s, t) - (ast + bs + ct + d) \geq 0 \right\} \leq 4N(-d).$$

In addition we obtain a formula for

$$P\{\sup_{\partial D} |X(s, t)| - (ast + bs + ct + d) \geq 0\}, \quad a, b, c, d \geq 0.$$

Some results on two-parameter Brownian bridge are also included.

2. Some lemmas. To avoid unnecessary repetitions in the proofs of the theorems, the following lemmas are given. Throughout this paper $W(t)$ and $X(s, t)$ will denote the standard Wiener process and the Yeh-Wiener process, respectively.

LEMMA 1. (Doob [2: p. 398]). *If $a \geq 0, b > 0, \alpha \geq 0, \beta > 0$, then*

$$\begin{aligned} & P \left\{ \sup_{0 \leq t < \infty} [W(t) - (at + b)] \geq 0 \quad \text{or} \quad \inf_{0 \leq t < \infty} [W(t) + at + \beta] \leq 0 \right\} \\ &= \sum_{m=1}^{\infty} \exp \{ -2[m^2 ab + (m - 1)^2 \alpha \beta + m(m - 1)(a\beta + \alpha b)] \} \\ & \quad + \exp \{ -2[(m - 1)^2 ab + m^2 \alpha \beta + m(m - 1)(a\beta + \alpha b)] \} \end{aligned}$$

$$\begin{aligned}
 & - \exp \{ -2[m^2(ab + \alpha\beta) + m(m - 1)\alpha\beta + m(m + 1)ab] \} \\
 & - \exp \{ -2[m^2(ab + \alpha\beta) + m(m - 1)a\beta + m(m - 1)ab] \} .
 \end{aligned}$$

LEMMA 2. *Let $f(t)$ be a Borel measurable function. Then for each Borel set E of real numbers,*

$$\begin{aligned}
 (2.1) \quad & P\{W(t) - f(t) \in E, 0 < t \leq 1 \mid W(1) = u\} \\
 & = P\left\{W(t) + u - (t + 1)f\left(\frac{1}{t+1}\right) \in \frac{1}{t}E, 0 < t < \infty\right\} .
 \end{aligned}$$

Proof. The basic technique used here is the same as the one used by Malmquist in [5]. Observe that $W(t)$ and $tW(1/t)$ are equivalent processes for $t > 0$. Thus, the left-hand side of (2.1) reduces to

$$\begin{aligned}
 & P\left\{W\left(\frac{1}{t}\right) - \frac{1}{t}f(t) \in \frac{1}{t}E, 0 < t \leq 1 \mid W(1) = u\right\} \\
 & = P\left\{W\left(\frac{1}{t}\right) - W(1) - \left[\frac{1}{t}f(t) - u\right] \right. \\
 & \quad \left. \in \frac{1}{t}E, 0 < t \leq 1 \mid W(1) = u\right\} .
 \end{aligned}$$

Upon using the fact that $W(1/t - 1)$ and $W(1/t) - W(1)$ are equivalent processes for $t > 0$, and $W(1/t) - W(1)$ and $W(1)$ are independent for $1 \geq t > 0$, we have the result by the transformation $1/t - 1 \rightarrow t$.

LEMMA 2.a. *If $f(t)$ is a Borel measurable function on $[0, 1]$, then*

$$\begin{aligned}
 & P\left\{\sup_{0 \leq t \leq 1} |X(1, t) - f(t)| \geq 0 \mid X(1, 1) = u\right\} \\
 & = P\left\{\sup_{0 \leq t < \infty} |X(1, t) + u| - (t + 1)f\left(\frac{1}{t+1}\right) \geq 0\right\} ,
 \end{aligned}$$

and the same holds for $X(t, 1)$.

LEMMA 3. *Let $f(s, t)$ be a Borel measurable function on D . Then for each Borel set E of real numbers,*

$$\begin{aligned}
 & P\{X(s, t) - f(s, t) \in E, (s, t) \in (0, 1]^2 \mid X(1, 1) = u\} \\
 & = P\left\{X(s + 1, t + 1) - X(1, 1) \right. \\
 & \quad \left. - \left[(s + 1)(t + 1)f\left(\frac{1}{s+1}, \frac{1}{t+1}\right) - u\right] \in \frac{E}{st}, (s, t) \in (0, \infty)^2\right\} .
 \end{aligned}$$

Proof. This lemma is a two-parameter analogue of Lemma 2, and it can be proved similarly by observing that $X(s, t)$ and $sX(1/s, 1/t)$ are equivalent processes for $s, t > 0$.

LEMMA 4. *Let $f(t)$ and $g(t)$ be any Borel measurable functions on $[0, 1]$. Then for any Borel sets E_1 and E_2 of real numbers,*

$$(2.2) \quad \begin{aligned} & P\{X(s, 1) - f(s) \in E_1, X(1, t) - g(t) \in E_2, (s, t) \in D \mid X(1, 1) = u\} \\ &= P\{X(s, 1) - f(s) \in E_1, 0 \leq s \leq 1 \mid X(1, 1) = u\} \\ &\quad \cdot P\{X(1, t) - g(t) \in E_2, 0 \leq t \leq 1 \mid X(1, 1) = u\}. \end{aligned}$$

Proof. Observe first that $X(s, 1)$ and $sX(1/s, 1)$ are equivalent standard Wiener processes for $s > 0$, and so are $X(1, t)$ and $tX(1, 1/t)$ for $t > 0$. Now $s[X(1/s, 1) - X(1, 1) + u]$ and $t[X(1, 1/t) - X(1, 1) + u]$ are independent processes for $1 \geq s, t > 0$, and they are also independent of $\{X(s, t): (s, t) \in D\}$. Hence (2.2) gives:

$$(2.3) \quad \begin{aligned} & P\{X(s, 1) - f(s) \in E_1, X(1, t) - g(t) \in E_2, (s, t) \in D \mid X(1, 1) = u\} \\ &= P\{s[X(1/s, 1) - X(1, 1) + u] - f(s) \in E_1, 0 < s \leq 1\} \\ &\quad \cdot P\{t[X(1, 1/t) - X(1, 1) + u] - g(t) \in E_2, 0 < t \leq 1\}. \end{aligned}$$

But the two probabilities on the right-hand side of (2.3) are equal to $P\{X(s, 1) - f(s) \in E_1, 0 \leq s \leq 1 \mid X(1, 1) = u\}$ and $P\{X(1, t) - g(t) \in E_2, 0 \leq t \leq 1 \mid X(1, 1) = u\}$ respectively, and hence the proof is complete.

3. Main results and proofs. In what follows $\{X(s, t): s, t \geq 0\}$ will be used exclusively for the Yeh-Wiener process.

THEOREM 1. *If $a, b, c, d \geq 0$, then with $\bar{a} = a + b + c + d$,*

$$\begin{aligned} & P\left\{\sup_{s, t} X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ &= N(-\bar{a}) + e^{-2(a+b)(c+d)}N(a + b - c - d) \\ &\quad + e^{-2(a+c)(b+d)}N(a - b + c - d) \\ &\quad - e^{2(d-a)(b+c+2d)}N(a - b - c - 3d). \end{aligned}$$

Proof. First observe that

$$(3.1) \quad \begin{aligned} P_1 &\equiv P\left\{\sup_{s, t} X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ &= P\left\{\sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0\right\} \\ &\quad + P\left\{\sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0\right\} \end{aligned}$$

$$- P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \right\} .$$

Since $X(s, 1)$ and $X(1, t)$ are equivalent to the standard Wiener process $W(t)$, the first two probabilities on the right of (3.1) can be evaluated explicitly.

Now,

$$\begin{aligned} P_2 &\equiv P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \right\} \\ &= P \{ X(1, 1) \geq \bar{a} \} \\ &\quad + \int_{-\infty}^{a+b+c+d} P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \right. \\ &\quad \left. \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \mid X(1, 1) = u \right\} dN(u) . \end{aligned}$$

Due to the fact that

$$\begin{aligned} &P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \mid X(1, 1) = u \right\} \\ (3.2) \quad &= P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0 \mid X(1, 1) = u \right\} \\ &\quad \cdot P \left\{ \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \mid X(1, 1) = u \right\} , \end{aligned}$$

we may use Lemma 2 to get

$$\begin{aligned} P_2 &= N(-\bar{a}) \\ &\quad + \int_{-\infty}^{a+b+c+d} P \left\{ \sup_{s \geq 0} X(s, 1) - [(c + d)s + (\bar{a} - u)] \geq 0 \right\} \\ &\quad \cdot P \left\{ \sup_{t \geq 0} X(1, t) - [(b + d)t + (\bar{a} - u)] \geq 0 \right\} dN(u) \\ &= N(-\bar{a}) \\ &\quad + \int_{-\infty}^{a+b+c+d} e^{-2(c+d)(\bar{a}-u)} e^{-2(b+d)(\bar{a}-u)} dN(u) \\ &= N(-\bar{a}) + e^{2(d-a)(b+c+2d)} N(a - b - c - 3d) . \end{aligned}$$

The result now readily follows.

COROLLARY. *If $d \geq 0$, then*

$$P\left\{\sup_{\partial D} X(s, t) \geq d\right\} = 3N(-d) - e^{4d^2}N(-3d).$$

This corollary agrees with the result in [6: p. 877].

THEOREM 2. *If $\{Y(s, t): (s, t) \in D\}$ is the two-parameter Brownian bridge, i.e., $\{Y(s, t): (s, t) \in D\} = \{X(s, t): (s, t) \in D | X(1, 1) = 0\}$ and $a, b, c, d \geq 0$, then*

$$P\left\{\sup_{\partial D} Y(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ = e^{-2(b+d)\bar{a}} + e^{-2(b+d)\bar{a}} - e^{-2(b+c+2d)\bar{a}}.$$

Proof. This follows from (3.2) by setting $u = 0$.

THEOREM 3. *If $a, b, c \geq 0$ and $d > 0$, then with $\bar{a} = a + b + c + d$ and $\bar{c} = c + d$,*

$$P\left\{\sup_{\partial D} \frac{|X(s, t)|}{ast + bs + ct + d} \geq 1\right\} = 2f(a, b, c, d),$$

where

$$f(a, b, c, d) = N(-\bar{a}) + \sum_{k=1}^{\infty} (-1)^{k+1} \left[e^{-2(a+b)\bar{c}k^2} \int_{-\bar{a}-2\bar{c}k}^{\bar{a}-2\bar{c}k} dN(u) \right. \\ \left. + e^{-2(a+c)(b+d)k^2} \int_{-\bar{a}-2(b+d)k}^{\bar{a}-2(b+d)k} dN(u) \right] \\ - \sum_{j,k=1}^{\infty} (-1)^{j+k} e^{-2\bar{a}[\bar{c}j^2+(b+d)k^2]} \left\{ e^{2[\bar{c}j+(b+d)k]^2} \right. \\ \left. \times \int_{-\bar{a}-2[\bar{c}j+(b+d)k]}^{\bar{a}-2[\bar{c}j+(b+d)k]} dN(u) + e^{2[\bar{c}j-(b+d)k]^2} \int_{-\bar{a}-2[\bar{c}j-(b+d)k]}^{\bar{a}-2[\bar{c}j-(b+d)k]} dN(u) \right\}.$$

Proof. Observe that

$$P_3(u) \equiv P\left\{\sup_{\partial D} \frac{|X(s, t)|}{ast + bs + ct + d} \geq 1 \mid X(1, 1) = u\right\} \\ = P\left\{\sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \text{ or} \right. \\ \left. \sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a + c)t + (b + d)} \geq 1 \mid X(1, 1) = u\right\}.$$

Upon applying Lemma 4, we obtain

$$P_3(u) = P\left\{\sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u\right\} \\ + P\left\{\sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a + c)t + (b + d)} \geq 1 \mid X(1, 1) = u\right\}$$

$$\begin{aligned}
 & - P\left\{ \sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & \cdot P\left\{ \sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a + c)t + (b + d)} \geq 1 \mid X(1, 1) = u \right\}.
 \end{aligned}$$

Due to Lemma 2.a., it follows

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & = P\left\{ \sup_{0 \leq s < \infty} \frac{|X(s, 1) + u|}{(c + d)s + \bar{a}} \geq 1 \right\} \\
 & = P\left\{ \sup_{0 \leq s < \infty} X(s, 1) - [(c + d)s + (\bar{a} - u)] \geq 0 \right. \\
 & \quad \left. \text{or } \inf_{0 \leq s < \infty} X(s, 1) + [(c + d)s + (\bar{a} + u)] \leq 0 \right\}.
 \end{aligned}$$

Lemma 1 applied to the last expression gives:

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq s < \infty} X(s, 1) - [(c + d)s(\bar{a} - u)] \geq 0 \right. \\
 & \quad \left. \text{or } \inf_{0 \leq s < \infty} X(s, 1) + [(c + d)s + (\bar{a} + u)] \leq 0 \right\} \\
 & = \sum_{m=1}^{\infty} \{ e^{-2\bar{a}(c+d)(2m-1)^2} [e^{2(c+d)(2m-1)u} + e^{-2(c+d)(2m-1)u}] \\
 & \quad - e^{-2\bar{a}(c+d)(2m)^2} [e^{2(c+d)(2m)u} + e^{-2(c+d)(2m)u}] \}.
 \end{aligned}$$

Therefore

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & = \sum_{j=1}^{\infty} (-1)^{j+1} e^{-2\bar{a}(c+d)j^2} [e^{2(c+d)ju} + e^{-2(c+d)ju}]
 \end{aligned}$$

and

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a + c)t + (b + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & = \sum_{k=1}^{\infty} (-1)^{k+1} e^{-2\bar{a}(b+d)k^2} [e^{2(b+d)ku} + e^{-2(b+d)ku}].
 \end{aligned}$$

Since $X(1, 1)$ is the standard normal random variable, the result now follows by:

$$\begin{aligned}
 & P\left\{ \sup_{\partial D} \frac{|X(s, t)|}{ast + bs + ct + d} \geq 1 \right\} \\
 & = P\{|X(1, 1)| \geq \bar{a}\} + \int_{-\bar{a}}^{\bar{a}} P_{\bar{a}}(u) dN(u)
 \end{aligned}$$

$$= 2N(-\bar{a}) + \int_{-\bar{a}}^{\bar{a}} P_s(u) dN(u) .$$

THEOREM 4. *If $a, b, c, d \geq 0$ and $\bar{a} = a + b + c + d, \bar{b} = b + c + d, \bar{c} = c + d$, then for $u < \bar{a}$,*

$$P_+ \equiv P\left\{ \sup_D X(s, t) - (ast + bs + ct + d) \geq 0 \mid X(1, 1) = u \right\} \\ \geq \begin{cases} \frac{\bar{b}}{b} [e^{2\bar{c}(u-\bar{a})} - e^{-2\bar{b}(u-\bar{a})}] + e^{2\bar{b}(u-\bar{a})}, & b > 0 \\ e^{2\bar{c}(u-\bar{a})} [1 + 2\bar{c}(\bar{a} - u)] & , \quad b = 0 . \end{cases}$$

Proof. Upon applying Lemma 3, we obtain

$$(3.3) \quad P_+ = P\left\{ \sup_{s, t \geq 0} X(s + 1, t + 1) - X(1, t + 1) + X(1, t + 1) - X(1, 1) \right. \\ \left. - [d(s + 1)(t + 1) + c(s + 1) + b(t + 1) + a - u] \geq 0 \right\} .$$

Consider the fact that $X(s + 1, t + 1) - X(1, t + 1)$ and $X(1, t + 1) - X(1, 1)$ are independent processes equivalent to $X(s, t + 1)$ and $X(1, t)$, respectively. The latter $X(1, t)$ will be denoted by $X^*(1, t)$ to signify that it is independent of $X(s, t + 1)$. Due to the fact that $c(s + 1) \leq c(s + 1)(t + 1)$ for all $c, s, t \geq 0$, it follows from (3.3)

$$(3.4) \quad P_+ \geq P\left\{ \sup_{s, t \geq 0} X(s, t + 1) + X^*(1, t) - [\bar{c}(t + 1)s + \bar{b}t + \bar{a} - u] \geq 0 \right\} \\ \geq \int_{u-\bar{a}}^{\infty} P\left\{ \sup_{s \geq 0} X(s, t + 1) - [\bar{c}(t + 1)s - r] \geq 0 \mid \sup_{t \geq 0} X^*(1, t) \right. \\ \left. - (\bar{b}t + \bar{a} - u) = r \right\} p(r, u) dr ,$$

where $p(r, u)$ is the probability density of

$$P\left\{ \sup_{t \geq 0} X^*(1, t) - (\bar{b}t + \bar{a} - u) \leq r \right\} \\ = \begin{cases} 1 - e^{-2\bar{b}(\bar{a} + r - u)}, & u - \bar{a} \leq r \\ 0 & , \quad \text{otherwise .} \end{cases}$$

Thus

$$(3.5) \quad p(r, u) = \begin{cases} 2\bar{b}e^{-2\bar{b}(\bar{a} + r - u)}, & u - \bar{a} \leq r \\ 0 & , \quad \text{otherwise .} \end{cases}$$

Observe that the probability in the integrand of (3.4) becomes

$$(3.6) \quad P \left\{ \sup_{s \geq 0} X(s, t + 1) - [\bar{c}(t + 1)s - r] \geq 0 \right\} \\ = \begin{cases} e^{2\bar{c}r}, & r \leq 0 \\ 1, & r > 0. \end{cases}$$

Therefore, (3.5) and (3.6) together with (3.4) give

$$P_4 \geq \int_{u-\bar{a}}^0 e^{2\bar{c}r} 2\bar{b}e^{-2\bar{b}(\bar{a}+r-u)} dr + \int_0^\infty 2\bar{b}e^{-2\bar{b}(\bar{a}+r-u)} dr,$$

from which the result readily follows.

The following is a special case ($u = 0$) of Theorem 4, which has broad application in Kolmogorov-Smirnov statistics.

THEOREM 4.a. *If $\{Y(s, t): (s, t) \in D\}$ is the two-parameter Brownian bridge and if $a, b, c, d \geq 0$, then*

$$P\{\sup_D Y(s, t) - (ast + bs + ct + d) \geq 0\} \\ \geq \begin{cases} \frac{\bar{b}}{b}(e^{-2\bar{a}\bar{a}} - e^{-2\bar{a}\bar{b}}) + e^{-2\bar{a}\bar{b}}, & b > 0 \\ (1 + 2\bar{a}\bar{c})e^{-2\bar{a}\bar{b}}, & b = 0. \end{cases}$$

THEOREM 5. *If $a, b, c, d \geq 0$, then*

$$P\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\} \\ \geq \begin{cases} N(-\bar{a}) + \frac{\bar{b}}{b}[N(\bar{a} - 2\bar{c})e^{-2\bar{c}(\bar{a}+b)} - N(\bar{a} - 2\bar{b})e^{-2\bar{a}\bar{b}}] \\ \quad + N(\bar{a} - 2\bar{b})e^{-2\bar{a}\bar{b}}, & b > 0 \\ N(-\bar{a}) + \frac{2\bar{c}}{\sqrt{2\pi}}e^{-\bar{a}^2/2} + N(\bar{a} - 2\bar{c})(1 + 2\bar{a}\bar{c} - 4\bar{c}^2)e^{-2\bar{a}\bar{c}}, & b = 0. \end{cases}$$

In particular,

$$P\{\sup_D X(s, t) - \lambda \geq 0\} \geq 2 \left[(1 - \lambda^2)N(-\lambda) + \frac{\lambda}{\sqrt{2\pi}}e^{-\lambda^2/2} \right] \\ = 2 \left[N(-\lambda) + \lambda \int_1^\infty N(-s)ds \right], \quad \lambda \geq 0.$$

Proof. The theorem now can be established by integrating lower estimates of the conditional probability P_4 in Theorem 4 with respect to $dP\{X(1, 1) \leq u\} = dN(u) = (2\pi)^{-1/2} \exp(-u^2/2)du$. The special case when $a = b = c = 0$ and $d = \lambda$ agrees with Goodman's result (Theorem 3 in [3]).

In order to find sharper upper bounds for the barrier-crossing

probabilities we introduce the following: Let $f(s, t)$ be a continuous function on D . If $\sup_D X(s, t) - f(s, t) \geq 0$, then define $\tau_f = (s_0, t_0)$ where

$$s_0 = \inf \{s \in [0, 1] \mid X(s, t) = f(s, t) \text{ for some } t \in [0, 1]\},$$

$$t_0 = \inf \{t \in [0, 1] \mid X(s_0, t) = f(s_0, t)\},$$

while if $\sup_D X(s, t) - f(s, t) < 0$, then set $\tau_f = (\infty, \infty)$. Thus with the convention that $(s_1, t_1) \leq (s_2, t_2)$ if and only if $s_1 \leq s_2$ and $t_1 \leq t_2$, we have that

$$P\left\{\sup_D X(s, t) - f(s, t) \geq 0\right\} = P\left\{\tau_f \leq (1, 1)\right\}.$$

THEOREM 6. *If $c, d \geq 0$, then*

$$P\left\{\sup_D X(s, t) - (ct + d) \geq 0\right\}$$

$$\leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\}$$

$$= 2[1 - N(c + d) + \exp(-2cd)N(a - b)].$$

Proof. Let τ stand for τ_f when $f(s, t) = ct + d$. Define

$$F(s, t) \equiv P\{\tau \leq (s, t)\}.$$

Then

$$F(1, 1) = P\left\{\sup_D X(s, t) - (ct + d) \geq 0\right\}$$

$$= P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\}$$

$$+ P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) < 0, \sup_D X(s, t) - (ct + d) \geq 0\right\}$$

$$(3.7) \quad = P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\}$$

$$+ \int_0^1 P\left\{\sup_{0 \leq t' \leq 1} X(1, t') - (ct' + d) < 0 \mid \tau = (s, t)\right\} dF(s, t)$$

$$\leq P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\}$$

$$+ \int_0^1 P\left\{X(1, t) - (ct + d) < 0 \mid \tau = (s, t)\right\} dF(s, t).$$

On account of the fact that $\tau = (s, t)$ implies $X(s, t) = ct + d$ and $X(1, t) - X(s, t)$ is independent of the conditioning $\tau = (s, t)$, it follows that

$$\begin{aligned}
 (3.8) \quad & \int_0^1 P\{X(1, t) - (ct + d) < 0 \mid \tau = (s, t)\} dF(st) \\
 & = \int_0^1 P\{X(1, t) - X(s, t) < 0\} dF(s, t) = \frac{1}{2} F(1, 1).
 \end{aligned}$$

The theorem now follows readily from (3.7) and (3.8).

COROLLARY 6.1. *If $b, c, d \geq 0$, then*

$$\begin{aligned}
 (3.9) \quad & P\left\{\sup_D X(s, t) - (bs + ct + d) \geq 0\right\} \\
 & \leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (b^*t + d) \geq 0\right\}, \quad b^* = \max\{b, c\}.
 \end{aligned}$$

Proof. The result follows immediately by observing that

$$\begin{aligned}
 & P\left\{\sup_D X(s, t) - (bs + ct + d) \geq 0\right\} \\
 & \leq \min\left\{P\left[\sup_D X(s, t) - (bs + d) \geq 0\right], \right. \\
 & \quad \left.P\left[\sup_D X(s, t) - (ct + d) \geq 0\right]\right\}.
 \end{aligned}$$

The right-hand side of (3.9) can also serve as an upper bound of $P\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\}$, and it is certainly a substantial improvement over (1.9). We state this fact formally as a corollary.

COROLLARY 6.2. *If $a, b, c, d \geq 0$, then*

$$\begin{aligned}
 (3.10) \quad & P\left\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\
 & \leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (b^*t + d) \geq 0\right\} \\
 & \leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - d \geq 0\right\} = 4N(-d),
 \end{aligned}$$

where $b^* = \max\{b, c\}$.

4. Supremum over rectangular regions. Some adjustments are needed to apply the results for the more general rectangular region $Q = [0, S] \times [0, T]$. The conversion formulas are given by:

$$\begin{aligned}
 (4.1) \quad & P\left\{\sup_{\partial Q} X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\
 & = P\left\{\sup_{\partial D} X(s, t) - (a'st + b's + c't + d') \geq 0\right\},
 \end{aligned}$$

where $a' = a\sqrt{ST}$, $b' = b\sqrt{S/T}$, $c' = c\sqrt{T/S}$, and $d' = d/\sqrt{ST}$.

$$(4.2) \quad P \left\{ \sup_q X(s, t) - (ast + bs + ct + d) \geq 0 \mid X(S, T) = u \right\} \\ = P \left\{ \sup_D X(s, t) - (a'st + b's + c't + d') \geq 0 \mid X(1, 1) = u' \right\},$$

where a', b', c', d' are as in (4.1) and $u' = u/\sqrt{ST}$. In (4.1), if ∂Q is replaced by Q , then D replaces ∂D .

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