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**THE EXPECTED MEASURE OF THE LEVEL SETS OF A
REGULAR STATIONARY GAUSSIAN PROCESS**

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If $X(t_1, t_2, \dots, t_d)$ is a sufficiently regular, centered, stationary Gaussian process, the (random) level set over a measurable domain $T \subset R^d$

$$A(u) = \{t \in T: X(t) = u\}$$

is a $d - 1$ -dimensional manifold embedded in R^d . Our main result states that its expected measure is given by

$$(1) \quad E\mu_{d-1}(A(u)) = \lambda(T)E \|\text{grad } X\| e^{-u^2/2}/\sqrt{2\pi}$$

where $\mu_{d-1}(A)$ is the $d - 1$ -dimensional volume of the hypersurface A , λ is the Lebesgue measure on R^d and the variance of X is assumed to be one.

The expression (1) holds even for $d = 1$. In that case $\mu_0(A)$ is a counting measure that gives the number of points in A . (μ_1 and μ_2 give respectively length and area.)

1. Preliminary notations and results.

DEFINITION 1. (i) A stationary centered Gaussian process X with parameter $t = (t_1, t_2, \dots, t_d) \in R^d$ and covariance function

$$(2) \quad \Gamma(t) = EX(s)X(s + t)$$

is said to be regular when it has continuous derivatives $\dot{X} = \text{grad } X = (X_1, \dots, X_d)$, $X_i(t) = \partial X(t)/\partial t_i$ ($i = 1, \dots, d$) and Γ is continuously differentiable up to the sixth order.

(ii) When, in addition, $\Gamma(0) = -\Gamma_{ii}(0) = 1$ ($i = 1, 2, \dots, d$) and $\Gamma_{ij}(0) = 0$ ($i \neq j$, $i, j = 1, \dots, d$), the process is said to be normalized. (Here and in what follows, the partial derivatives of Γ are denoted by $\Gamma_i = \partial\Gamma/\partial t_i$, $\Gamma_{ij} = \partial^2\Gamma/\partial t_i\partial t_j$, \dots .)

The strong requirement imposed to the covariance of a regular process (which is justified by the use of a theorem by R. J. Adler and A. M. Hasofer, here stated as Lemma 1 (ii)) largely implies the existence of a version of \dot{X} with continuous derivatives (see for instance [5]). The vector variable

$$\dot{X} = (X_1, X_2, \dots, X_d)$$

has covariances

$$\text{Var } \dot{X} = -((\Gamma_{ij}(0))) .$$

A change of scale in the process and a linear change of para-

meter, lead to a normalized process, namely,

$$Y(t) = X\left(\left(\left(\frac{-\Gamma_{ij}(0)}{\Gamma(0)}\right)\right)^{-1/2} t\right) / \sqrt{\Gamma(0)}.$$

Even if a process is not normalized, we shall assume in what follows, without loss of generality, that $\Gamma(0) = 1$.

The following lemma states known results. We indicate the corresponding references and omit the proofs.

LEMMA 1. *Let X be a regular process. Then:*

(i) (Belyaiev [3] Thm. 3.2). *Given u and the interval*

$$(3) \quad T = \prod_{i=1}^d [a_i, b_i] \subset R^d,$$

with probability one there is no point $t \in T$ such that

$$(4) \quad X(t) = u \quad \text{and} \quad \dot{X}(t) = 0.$$

(ii) (Adler [1] (proof of Theorem 2) and Hasofer [2, 7]). *The number of points $t \in T$ such that $X(t) = u$ and all but one of the scalar conditions (4) hold, has a finite expectation. (The references give in fact the actual value of the expectation.)*

Given T by (3), let us introduce the notations $T_i^{(-1)} = \{a_i\}$, $T_i^{(0)} = (a_i, b_i)$, $T_i^{(1)} = \{b_i\}$. If $k = (k_1, \dots, k_d)$ is a multi-index with components $k_i = -1, 0$ or 1 ($i = 1, \dots, d$), we abbreviate T^k for $\prod_{i=1}^d T_i^{(k_i)}$. The set T^k will be called a face of T of dimension $|k| = \sum_{i=1}^d (1 - |k_i|)$. (In particular, the interior T^0 of T is the only d -dimensional face.)

DEFINITION 2. Given a $d - 1$ -dimensional manifold $A \subset T$ with continuous normal $\xi(t) = (\xi_1(t), \dots, \xi_d(t)) \neq 0$, $t \in A$, a point

$$t \in T^k \cap A$$

such that all but one of the $|k|$ conditions

$$\{\xi_i(t) = 0\}_{k_i=0}$$

hold, will be said to be a k -critical point.

COROLLARY 1. *The level set*

$$(5) \quad A(u) = \{t \in T: X(t) = u\}$$

of a regular process X is a $d - 1$ -manifold with continuous normal $\dot{X} \neq 0$ with probability one. The number $X_k^{(u)}$ of k -critical points

of $A(u)$ has a finite expectation for each $k \in K = \{-1, 0, 1\}^d$.

Proof. The first assertion follows readily from Definition 1 and Lemma 1(i). The second one follows from Lemma 1(ii) applied to the restriction of X to T^k , when $|k| > 1$.

The critical points for $|k| = 1$ are the crossings of the level by the one-dimensional restriction, and its (*finite*) expected number is computed in [5] by a well known formula, namely, (1) with $d = 1$. Finally, for $|k| = 0$, the conclusion is trivial.

LEMMA 2. *Given A as in Definition 2, if*

$$\psi_i = \psi_i(T) = \max_{t \in T} \# \{x \in R: t + xe_i \in A\},$$

$$e_i = (e_{i1}, e_{i2}, \dots, e_{id}), \quad e_{ij} = \begin{cases} 0 & \text{if } i \neq j \\ 1 & \text{if } i = j \end{cases},$$

and if \mathcal{L}_k is the number of k -critical points, then, for each $i = 1, \dots, d$,

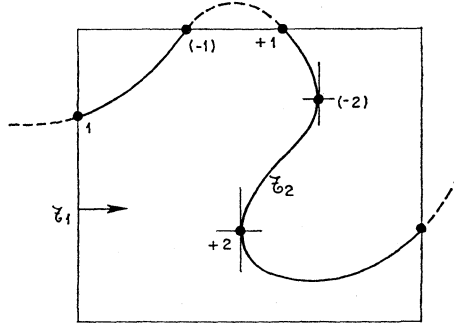
$$(6) \quad \Psi_i \leq 2 \sum_{k \in K} \mathcal{L}_k.$$

(Notice that Ψ_i is the maximum number of intersections with A of a line, parallel to the i th coordinate direction. Our estimate is a very rough one, but sufficient for our purposes.)

Proof. Let us assume $i = 1$ for ease of description, and proceed to sweep T starting with the closure of a one-dimensional face in the given direction, say $\mathcal{I}_1 = \bar{T}^{(0,1,1,\dots,1)}$. As a first step, let us translate this face in the second direction, until it describes the closure of the two-dimensional face $\mathcal{I}_2 = \bar{T}^{(0,0,1,1,\dots,1)}$. Then \mathcal{I}_2 is translated in the third direction until it describes $\mathcal{I}_3 = \bar{T}^{(0,0,0,1,\dots,1)}$ and this procedure is continued until $\mathcal{I}_d = \bar{T}^0 = T$ has been described.

At each step, the maximum number of intersections with A of the lines in the given direction already described, is estimated as follows: At the beginning, we count the intersections of \mathcal{I}_1 , which are precisely the critical points on the faces that compose \mathcal{I}_1 , namely $\mathcal{L}_{(1,1,\dots,1)} + \mathcal{L}_{(0,1,1,\dots)} + \mathcal{L}_{(-1,1,1,\dots)}$. When \mathcal{I}_1 is translated, each increase of the number of intersections (in the amount of one or two) is produced when the face passes through a critical point. This is a necessary condition, through a critical point may produce a decrease (of one or two intersections) or no change. Therefore $2 \sum_{T_k \subset \mathcal{I}_2} \mathcal{L}_k$ is an upper bound of $\Psi_i(\mathcal{I}_2)$. Now, \mathcal{I}_2 is translated, and each increase in the number of intersections must be produced

when the face transverses a critical point, hence, since the increase is in the amount of one or two as before, $\Psi_i(\mathcal{T}_3) \leq 2 \sum_{\tau_k \in \mathcal{T}_3} \mathcal{H}_k$. Going on in the same way, we reach finally the required inequality (6).



COROLLARY 2. *The maximum number $\Psi_i^{(u)}$ of intersections of lines in the i th direction with the level set $A(u)$ of a regular process X , has a finite expectation.*

Proof. Use Lemma 2 and Corollary 1.

2. The expected measure of $A(u)$. Given a regular process X , let T and $A(u)$ be defined by (3) and (5). We introduce the cones

$\mathcal{E}_i^\alpha \subset \mathbb{R}^d$ defined by

$$\mathcal{E}_i^\alpha = \{\xi = (\xi_1, \dots, \xi_d) : |\xi_j| < \alpha |\xi_i| \text{ for each } j \neq i\}$$

and denote their relative solid angle by

$$\nu(\mathcal{E}_i^\alpha) = \frac{\mu_{d-1}(\mathcal{E}_i^\alpha \cap \{\xi : \|\xi\| = 1\})}{\mu_{d-1}(\{\xi : \|\xi\| = 1\})}.$$

Since $|X_i| > 0$ on $A_i^\alpha(u) = \{t \in A(u) : \dot{X}(t) \in \mathcal{E}_i^\alpha\}$, the portion $A_i^\alpha(u)$ of $A(u)$ can be locally parametrized in the form $t_i = F(t_1, \dots, t_{i-1}, t_{i+1}, \dots, t_d)$, where F satisfies $\partial F / \partial t_j = -X_j / X_i$ ($j \neq i$) because of the Implicit Function Theorem.

Hence the $d - 1$ -dimensional volume of $A_i^\alpha(u)$ is given by the integral (see for instance [6], p. 334):

$$\mu_{d-1}(A_i^\alpha(u)) = \int_{t \in A_i^\alpha(u)} \frac{\|\dot{X}(t)\|}{|X_i(t)|} dt_1 \cdots dt_{i-1} dt_{i+1} \cdots dt_d.$$

If $\mathcal{D}_n = \{t \in T; 2^n t_i \text{ is an integer for each } i = 1, \dots, d\}$, is the set of diadic points in T , then the integral is approximated by the sums.

$$S_i^\alpha(n) = \sum_{\substack{t \in \mathcal{D}_n \\ t + 2^{-n} e_i \in T}} 2^{-n(d-1)} \mathbf{1}_{(-\infty, 0)}((x(t) - u)(x(t + 2^{-n} e_i) - u)) \mathbf{1}_{\mathcal{E}_i^\alpha}(\dot{X}(t)) \frac{\|\dot{X}(t)\|}{|X_i(t)|}.$$

Where 1_C is the indicator function of C , for any set C . More precisely, it is easily seen that for each $\varepsilon > 0$, $\lim_{n \rightarrow \infty} S_i^{\alpha - \varepsilon}(n) \leq \mu_{d-1}(A_i^\alpha(u)) \leq \lim_{n \rightarrow \infty} S_i^\alpha(n)$, and, since

$$1_{C_i^\alpha}(\dot{X}(t)) \frac{\|\dot{X}(t)\|}{|\dot{X}_i(t)|} \leq \sqrt{d} \alpha$$

and

$$\sum_{\substack{t \in \mathcal{I} \\ t + 2^{-n} e_i^n \in T}} 2^{-n(d-1)} 1_{(-\infty, 0)}((X(t) - u)(X(t + 2^{-n} e_i) - u)) \leq \Psi_i(u) \frac{\lambda(T)}{b_i - a_i},$$

then $S_i^\alpha(n)$ is dominated by the random variable $\sqrt{d} \alpha \Psi_i(u) (\lambda(T) / (b_i - a_i))$, whose expectation is finite from Corollary 2.

Applying the Dominated Convergence Theorem, and noticing that $\mu_{d-1}(A_i^\alpha(u))$ is increasing in α , it follows

$$(7) \quad \lim_{n \rightarrow \infty} ES_i^{\alpha - \varepsilon}(n) \leq E \mu_{d-1}(A_i^\alpha(u)) \leq \lim_{n \rightarrow \infty} ES_i^\alpha(n).$$

In order to compute $\lim_{n \rightarrow \infty} ES_i^\alpha(n)$, we write

$$\begin{aligned} ES_i^\alpha(n) &= \sum_{t \in \mathcal{I}} 2^{-nd} E \left(1_{C_i^\alpha}(\dot{X}(t)) \frac{2^n \|\dot{X}(t)\|}{|\dot{X}_i(t)|} P\{(X(t) - u)(X(t + 2^{-n} e_i) - u) \right. \\ &\quad \left. < 0 / \dot{X}(t)\} \right) \end{aligned}$$

and the stationarity of X leads to

$$\begin{aligned} \lim_{n \rightarrow \infty} ES_i^\alpha(n) &= \lambda(T) \lim_{n \rightarrow \infty} 2^n E \left(1_{C_i^\alpha}(\dot{X}(0)) \frac{\|\dot{X}(0)\|}{|\dot{X}_i(0)|} P\{(X(0) - u)(X(2^{-n} e_i) - u) \right. \\ &\quad \left. < 0 / \dot{X}(0)\} \right). \end{aligned}$$

Let us abbreviate $\Gamma_\cdot(\delta) = (\Gamma_1(\delta), \dots, \Gamma_d(\delta))^{tr}$. The conditional distribution of $(X(0), X(2^{-n} e_i))$ given $\dot{X}(0)$ is Gaussian, with expectation $(0, -\Gamma_\cdot(2^{-n} e_i)((-\Gamma_{h_j}(0)))^{-1} \dot{X}(0))$ and variance

$$\begin{pmatrix} 1 & \Gamma(\delta) \\ \Gamma(\delta) & 1 - \Gamma_\cdot^{tr}(2^{-n} e_i)((-\Gamma_{h_j}(0)))^{-1} \Gamma_\cdot(2^{-n} e_i) \end{pmatrix},$$

and the Taylor expansions of Γ , Γ_i are

$$\begin{aligned} \Gamma(t) &= 1 + \frac{1}{2} \sum_{i_j=1}^d \Gamma_{i_j}(0) t_i t_j + \frac{1}{24} \sum_{i_j k l=1}^d \Gamma_{i_j k l} t_i t_j t_k t_l + \dots \\ \Gamma_i(t) &= \sum_{j=1}^d \Gamma_{i_j}(0) t_j + \frac{1}{6} \sum_{j k l=1}^d \Gamma_{i_j k l} t_j t_k t_l + \dots, \end{aligned}$$

therefore, it is easily seen that the conditional distribution of

$$\left(X(0), Z_n = \frac{X(2^{-n}e_i) - X(0)}{2^{-n}\dot{x}_i} \right)$$

has conditional expectation of $(0, 1 + 0(2^{-n}))$ and conditional variance

$$\begin{pmatrix} 1 & 0(2^{-n}) \\ 0(2^{-n}) & 0(2^{-2n}) \end{pmatrix}.$$

Then we have

$$\begin{aligned} \lim_{n \rightarrow \infty} - \frac{2^n \|\dot{x}\|}{|\dot{x}_i|} P\{(X(0) - u)(X(2^{-n}e_i) - u) < 0/\dot{X}(0) = \dot{x}\} \\ = \lim_{n \rightarrow \infty} \frac{2^n \|\dot{x}\|}{|\dot{x}_i|} P\{(X(0) - u)(2^{-n}\dot{x}_i Z_n + X(0) - u) < 0/\dot{X}(0) = \dot{x}\} \\ = \frac{\|\dot{x}\|}{|\dot{x}_i|} \cdot |\dot{x}_i| \frac{1}{\sqrt{2\pi}} e^{-u^2/2} = \|\dot{x}\| \varphi(u), \end{aligned}$$

thus

$$\lim_{n \rightarrow \infty} ES_i^\alpha(n) = \lambda(T)\varphi(u)E(\mathbf{1}_{\varphi_i^\alpha(\dot{X})} \|\dot{X}\|).$$

This limit is a continuous function of α , hence

$$E\mu_{d-1}(A_i^\alpha(u)) = \lambda(T)\varphi(u)E(\mathbf{1}_{\varphi_i^\alpha(\dot{X})} \|\dot{X}\|);$$

furthermore, the inclusions

$$\bigcup_{i=1}^d A_i^1(u) \subset A(u) \subset \bigcup_{i=1}^d A_i^\alpha(u) \quad (\alpha > 1)$$

and the fact that $A_1^1(u), A_2^1(u), \dots, A_d^1(u)$ are disjoint, imply

$$\sum_{i=1}^d E\mu_{d-1}(A_i^1(u)) \leq E\mu_{d-1}(A(u)) \leq \sum_{i=1}^d E\mu_{d-1}(A_i^\alpha(u)).$$

We use again the continuity in α to obtain the result stated as follows.

THEOREM. (i) *The expected measure of the level set of a regular process X corresponding to a measurable set T and a level u , is*

$$(8) \quad E\mu_{d-1}(A(u)) = \lambda(T)\varphi(u)E\|\dot{X}\|,$$

where \dot{X} is Gaussian, $E\dot{X} = 0$, $\text{Var } \dot{X} = ((-\Gamma_{ij}(0)))$, and $\varphi(u) = 1/\sqrt{2\pi} e^{-u^2/2}$.

(ii) *When X is normalized, (8) reduces to*

$$E\mu_{d-1}(A(u)) = \lambda(T)e^{-u^{2/2}}/B\left(\frac{1}{2}, \frac{d}{2}\right)$$

where B is Euler Beta function.

The proof of (8) for an interval T is contained in the preceding context; since the expectation is an additive function of T , the same result holds for measurable T .

When X is normalized, a straightforward calculation gives the final result.

3. Comparisons with previous results. For $d = 1$, (8) reduces to the formula

$$\text{expected number of crossings of } u = \lambda(T)\sqrt{-\Gamma''(0)}e^{-u^2/2}/\pi,$$

given in [5].

In the case $d = 2$, Benzaquen [4] proved that if π_i is the projection $\pi_i(t_1, \dots, t_d) = (t_1, \dots, t_{i-1}, t_{i+1}, \dots, t_d)$ and $\mu_{d-1}^{(i)}(A(u)) = \mu_{d-1}(\pi_i(A(u)))$, where the points are taken with its corresponding multiplicity, then

$$(9) \quad E\mu_{d-1}^{(i)}(A(u)) \leq \lambda(T)\sqrt{-\Gamma_{ii}(0)}e^{-u^2/2}/\pi.$$

It is not hard to prove the equality in (9) with our assumptions of regularity, and to extend the same formula for $d > 2$.

Clearly, the inequalities

$$\mu_{d-1}^{(i)}(A(u)) \leq \mu_{d-1}(A(u)) \leq \sum_{j=1}^d \mu_{d-1}^{(j)}(A(u))$$

hold, and the compatibility of (8) and (9) require

$$\sqrt{2/\pi}\sqrt{-\Gamma_{ii}(0)} \leq E\|\dot{X}\| \leq \sqrt{2/\pi} \sum_{j=1}^d \sqrt{-\Gamma_{jj}(0)}$$

and, in the normalized case,

$$\sqrt{2/\pi} \leq \sqrt{2\pi}/B\left(\frac{1}{2}, \frac{d}{2}\right) \leq d\sqrt{2/\pi}.$$

These inequalities are trivially verified by a direct calculation of expectations in $|\dot{X}_i| \leq \|\dot{X}\| \leq \sum_{j=1}^d |\dot{X}_j|$.

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